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MISCELLANEOUS PAPER CERC-87-12



DURATION OF EXTREME WAVE CONDITIONS

by

Orson P. Smith

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Coastal Engineering Research Center

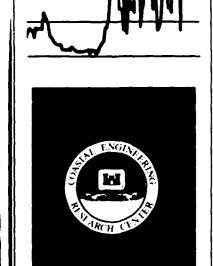
DEPARTMENT OF THE ARMY Waterways Experiment Station, Corps of Engineers PO Box 631, Vicksburg, Mississippi 39180-0631





June 1987 Final Report

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Prepared for DEPARTMENT OF THE ARMY US Army Corps of Engineers Washington, DC 20314-1000

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Statistical trends of the duration of extreme wave conditions, as characterized by hindcast wave information, are investigated at five sites along the coastline of the United States (three on the Atlantic coast and two on the Pacific coast). A review of pertinent statistical concepts and water wave characterization conventions and terminology is followed by a description of the Wave Information Studies Program of the US Army Engineer Waterways Experiment Station, Vicksburg, Miss. The database of hindcast wave information in shallow water created by this program is applied to develop a method of identification of extreme events and definition of their duration, based on exceedance of a threshold for zero moment wave heights. The number of events identified is found to be proportional to the percent exceedance of the specified threshold, regardless of geographical location. The Extremal Type I distribution is found to be superior to the Weibull distribution as a model for both distribution of durations and peak zero moment wave (Continued) 20 DISTRIBUTION/AVAILABILITY OF ABSTRACT OTIC USERS Unclassified Uncl								
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PREFACE

Work leading to preparation of this manuscript was conducted as part of the "Develop Functional and Structural Design Criteria" work unit of the Coastal Structures Evaluation and Design research and development program of the US Army Corps of Engineers. Authorization from the Office, Chief of Engineers, to publish this report is gratefully acknowledged. This report was originally submitted by the author to Mississippi State University as partial fulfillment of the requirements for an M.S. degree in Civil Engineering.

This study was conducted and the report written by Mr. Orson P. Smith, Coastal Design Branch (CDB), Wave Dynamics Division (WDD), of the Coastal Engineering Research Center (CERC), US Army Engineer Waterways Experiment Station (WES). The author acknowledges the assistance of Dr. Robert E. Jensen of the Coastal Oceanography Branch, Research Division, CERC, whose technical guidance with regard to application of the Wave Information Studies database of hindcast wave information was appreciated throughout the course of the work. The author also is grateful for the technical guidance of Dr. Michael E. Andrew of the Prototype Measurement and Analysis Branch, Engineering Development Division, CERC, regarding statistical aspects of the work.

Work was performed under the general direction of Dr. Frederick E. Camfield, Chief, CDB; Mr. C. E. Chatham, Chief, WDD; Mr. Charles C. Calhoun, Jr. Assistant Chief, CERC; and Dr. James R. Houston, Chief, CERC.

The Director of WES during the course of the work was COL Allen F. Grum, USA. Commander and Director of WES during publication of this report was COL Dwayne G. Lee, CE. Technical Director was Dr. Robert W. Whalin.

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CHAPTER I: INTRODUCTION

Statement of the Problem

Extreme wave conditions have been the cause of property loss, suffering, injuries, and death since man first approached the sea. Coastal engineers therefore have always attempted to build works that would withstand, with little or no damage, the worst impact of waves from very rare events. The direct effect of sea waves striking coastal structures has long been recognized as a critical phenomenon with respect to structural integrity during a storm at sea. The hydraulic impact of individual waves has traditionally been the specific force used as the basis of structural design criteria; therefore, characteristics of the worst few waves of a hypothetical extreme event have been estimated for application in most design computations. Rubble-mound structures, constructed of layered quarrystone or concrete shapes and built for centuries as wave barriers (breakwaters and jetties) or shore protection (revetments), are usually designed in this fashion.

The limits of functional performance of coastal structures have recently become more critical with respect to overall economic optimization. Public tinancing of coastal works has been more difficult to arrange than in past decades. The concept of designing a structure to be stable during a very extreme storm, but to be less than 100 percent effective in some extreme events of lesser intensity, has been in the minds of coastal engineers in an effort to conceive affordable harbor or shore protection plans. Life cycle cost also is receiving much more scrutiny, particularly with respect to expensive mobilization and challenging construction techniques required for repairs at many coastal projects. The bulwarks of extreme conservatism in coastal engineering design practice are beginning to buckle under pressure for more precise estimates of structural integrity and functional performance. These estimates may someday approach the precision of those now required for design of buildings and bridges.

One critical question in many new optimized designs is "What is the effect of duration of exposure?" Sandy beaches commonly change their shapes to a more stable configuration, given sufficient exposure to severe wave conditions, in theory approaching a new equilibrium (Bruun 1954). Some radical new rubble-mound concepts attempt to emulate this effect (Delft Hydraulics

Laboratory 1985). Laboratory experiments which simulate natural irregular waves also have shown some duration effects on rubble mounds of more traditional design (Graveson et al. 1980; Van der Meer and Pilarczyck 1984; and Tenaud et al. 1981). The open literature contains little specific guidance, however, for researchers or designers to estimate the duration of a given intensity of extreme wave conditions.

Purpose and Objectives

The purpose of this work is to investigate the duration of extreme wave conditions estimated from hindcast wave data, with a view toward developing a means to characterize the variation of these durations for use in design of coastal structures. Hindcast wave data, which are discussed later in more detail, are one of the most valuable tools of coastal engineers, primarily because weather data on which they are based typically exist for much longer periods of record than other wave information sources. The 20-year (1956-1975) Wave Information Studies (WIS) database of hindcast wave data prepared and maintained by the US Army Engineer Waterways Experiment Station (WES) (Brooks and Corson 1984) is a key source of wave information in many US Army Corps of Engineers projects since it now extends along most of the coastline of the United States.

The specific objectives of this study were to (1) review existing literature regarding the duration of extreme wave conditions and related topics; (2) formulate a practical means of identifying individual events of extreme wave conditions, relying on the intensity of wave conditions as represented in the WIS database and associated publications; (3) address the probability distribution of extreme event durations by fitting selected distribution functions to representative data; and (4) address the possible relation of an extreme event's duration to the peak conditions during the extreme event by regression analysis.

Organization

This report presents reviews of pertinent statistical concepts and techniques, considerations regarding the characterization of wave conditions, and the specific nature of WIS hindcast data before proceeding to describe the

progress toward and conclusion of the four objectives stated above. An overall summary and statement of conclusions then is followed by Appendix A containing figures and tables which were not presented in the main text for the sake of continuity and space conservation. Appendix B includes pertinent wave information transcribed from the WIS database. Appendix C includes a listing of the computer program STRMDIST which was used to identify extreme events, define durations, and fit parameterized distribution functions to both the durations and peak wave heights of extreme events identified. Appendix D includes the command file for the commercial statistical software package SPSS (Nie et al. 1975), which was applied to address the relationship of extreme event duration to peak wave conditions.

CHAPTER II: REVIEW OF PERTINENT STATISTICAL CONCEPTS

Continuous Frequency Distributions

The primary tools of this study are statistical procedures which address the variability of parameters of interest, specifically duration of extreme events at sea and their peak intensity. A brief review of pertinent statistical concepts, which are critical to understanding the methods and conclusions of the analysis, is presented below.

Continuous random variables are variables whose values are measured on a continuous scale, as opposed to their discrete counterparts such as rolling dice or coin flipping. Most natural phenomena of varying intensity as measured by instruments are treated as continuous random variables. The probability that the value of a particular random variable, x, will fall within a certain range can be estimated by application of its probability density function, f(x), which is analogous to a histogram for discrete variables. The following two conditions apply in defining probability density functions:

 $f(x) \ge 0$ for all x within the domain of f

and

$$\int_{-\infty}^{\infty} f(x) dx = 1$$
 (1)

The probability that x will fall within the range from a to b is given by:

$$P(a \le x \le b) = \int_{a}^{b} f(x) dx$$
 (2)

Technically the probability of x taking on a value of exactly a or b is zero, but since physical measurements cannot be infinitely accurate, the interval from a to b can be considered inclusive. A transformation of the

probability density function into its corresponding distribution function, F(x), allows more expedient computation of probabilities:

$$F(x) = \int_{-\infty}^{x} f(t) dt$$
 (3)

where f(t) is the probability density function of a dummy variable t.

The value of F(x) varies between 0 and 1. The probability that x will have a value equal to or less than a is F(a). The probability that x will have a value between a and b is F(b) - F(a). The corresponding probability density function is:

$$f(x) = \frac{dF(x)}{dx} \tag{4}$$

It is important to define the domain of f and that this domain include all the values of x of interest. Furthermore, the function f must be integrable within this domain (and F differentiable) for the above definitions to apply (Miller and Freund 1985).

Distribution Parameters

The mean or expected value of x is defined by:

$$\mu = \int_{-\infty}^{\infty} xf(x) dx$$
 (5)

The variance of probability density function is the expected value of the squared deviation from the mean, given by:

$$\sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$$
 (6)

The variance, σ^2 , and its square root, the standard deviation, σ , are both measures of the spread of the probability density about the mean. The standard deviation is expressed in the same units as x and μ . A small variance or standard deviation implies a strong central tendency while large values imply significant spread or "variance" of x values (Miller and Freund 1985).

The Poisson Distribution

A wide variety of distribution functions have been formulated by researchers and statisticians which have been shown to describe well the behavior of certain random variables which occur in nature. One such function is the Poisson distribution, defined by:

$$f(x) = \frac{\lambda^{x}e^{-\lambda}}{x!}$$
 for $x = 0, 1, 2, ...$ (7)

This is a discrete distribution which has important associations with the continuous distributions that have been applied to describe weather-related variables. Specifically, the roisson distribution has been applied to describe the number of occurrences of events taking place randomly over continuous intervals of time. The parameter λ is both the mean and the variance of the Poisson distribution. A key assumption behind application of this distribution is that the probability of an occurrence for the type of event in question during a small interval of time must not depend on what happened prior to that time. A random process which fits this criterion is called a Poisson process.

The Exponential Distribution

A continuous distribution which is often associated with the Poisson distribution is the exponential distribution, given by:

$$f(x) = \frac{e^{-x/\beta}}{\beta} \qquad \text{for } x > 0 \quad \text{and} \quad \beta > 0$$
$$= 0 \quad \text{elsewhere}$$
 (8)

The corresponding distribution function is:

$$F(x) = 1 - e^{-x/\beta} \tag{9}$$

The mean and standard deviation of a variable represented by an exponential distribution are both β and the variance is β^2 . This distribution is often used with Poisson processes to model the waiting time between successive occurrences. If the λ parameter of a Poisson distribution is the average number of occurrences in time T, then the average rate of occurrences per unit time is λ/T . The corresponding exponential distribution parameter is β = T/λ . This relation and the fact that both distributions are fully described by a single parameter make them easy to use in a wide range of applications dealing with the frequency of and waiting time between discrete events.

The Weibull Distribution

Another distribution, which is widely used to model the variation in intensities of natural extremes such as flood elevations and storm intensities, is the Weibull distribution, where:

$$f(x) = \frac{1}{\beta^{\alpha}} \alpha x^{\alpha - 1} \exp \left[-\left(\frac{x}{\beta}\right)^{\alpha} \right] \quad \text{for } x > 0 , \quad \alpha > 0 , \quad \beta > 0$$

$$= 0 \quad \text{elsewhere}$$
(10)

The corresponding Weibull distrubution function is very similar to the exponential distribution:

$$F(x) = 1 - \exp \left[-\left(\frac{x}{\beta}\right)^{\alpha} \right]$$
 (11)

The parameter $_{\alpha}$ is the "shape parameter" which defines the basic shape of the function. The $_{\beta}$ parameter is the "scale parameter" which determines the degree of spread along the abscissa (Isaacson and MacKensie 1981). The mean and variance of the Weibull distribution are:

$$\mu = \beta \Gamma \left(1 - \frac{1}{\alpha} \right) \tag{12}$$

$$\sigma^2 = \beta^2 \left[\Gamma \left(1 + \frac{2}{\alpha} \right) - \Gamma^2 \left(1 + \frac{1}{a} \right) \right]$$
 (13)

The gamma function is given by:

$$\Gamma(z) = \int_{0}^{\infty} x^{z-1} e^{-x} dx = (z-1)!$$
 (14)

The Weibull distribution has two parameters which make it actually a family of functions. A three-parameter form is sometimes used to provide further flexibility in adapting the distribution to certain phenomena, where:

$$F(x) = 1 - \exp\left[-\frac{(x - \varepsilon)}{\beta}\right]^{\alpha} \quad \text{for } \varepsilon > 0$$
 (15)

The parameter $^{\epsilon}$ is a "location parameter" which locates the position of the probability along the abscissa (x-axis). In the particular case of the Weibull distribution, $^{\epsilon}$ is in effect a lower limit to values of x. The $^{\epsilon}$ parameter is often taken as zero in practice. The Weibull distribution reduces to the exponential distribution when $^{\alpha}$ = 1 and $^{\epsilon}$ = 0 (Isaacson and MacKensie 1981).

The Rayleigh Distribution

The Weibull distribution reduces to a Rayleigh distribution when $\alpha=2$ and $\epsilon=0$, a function widely used to model the distribution of wave heights passing a point during a stationary sea state. The term "stationary" refers to the common assumption that, for practical purposes, statistical properties of ocean waves tend to be time invariant during a period of a few minutes to an hour or more. The time for significant changes to occur in a sea state is thus assumed to be substantially longer than the time necessary to measure the form of a few hundred waves passing a fixed point. The Rayleigh distribution, for this purpose, is often expressed in the form:

$$-2(H/H_g)^2$$
F(H) = 1 - e (16)

where H is an individual wave height in a sea state and H_S is the "significant wave height," also defined as the average of the highest 1/3 waves. This relation has been found to be quite accurate in most conditions at sea, with the exception of waves nearing the point of breaking in shallow water (Massie 1976). The corresponding probability density function, mean, and variance of this form of the Rayleigh distribution are:

$$f(x) = 4\left(\frac{H}{H_g^2}\right) e^{-2(H/H_g)^2}$$
 (17)

$$\mu = \left(\frac{\pi}{8}\right)^{1/2} H_{s} = 0.627 H_{s} \tag{18}$$

$$\alpha^2 = \left(\frac{1 - \pi}{8}\right) H_g^2 \qquad (\sigma = 0.779 H_g)$$
 (19)

The Extremal Type I Distribution

This distribution, sometimes called the "Gumbel" or "Fisher-Tippet Type I" distribution, also is frequently applied to model natural extremes such as storm intensities (Gumbel 1958). The probability density and distribution functions have the following forms:

$$f(x) = \frac{e^{-e^{-[(x-\epsilon)/\phi]}}}{\beta} \qquad \text{for } -\infty < x < \infty \\ -\infty < \epsilon < \infty \\ \beta > 0$$
 (20)

$$F(x) = e^{-e^{-[(x-\epsilon)/\beta]}}$$
(21)

The mean and variance are:

$$\mu = \varepsilon - \gamma \beta \tag{22}$$

$$\sigma^2 = \frac{\pi^2 \beta^2}{6} \tag{23}$$

where γ = Euler's constant = 0.5772 . The Extremal Type I distribution is also a two-parameter family of functions, in this case with a shape parameter of α = 1 in keeping with the usual practice for application to weather-related phenomena (Isaacson and MacKensie 1981 and Andrew et al. 1985). The ϵ parameter is again the location parameter and β the scale parameter. The Extremal Type I distribution is not constrained to positive values of x.

Figure 1 illustrates the relative form of the Exponential, Weibull, Rayleigh, and Extremal Type I distributions. The Exponential and Rayleigh

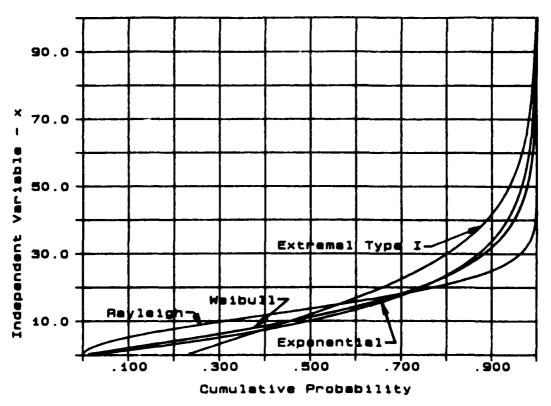


Figure 1. Relative form of four distribution functions curves shown in Figure 1 have the same mean as the Weibull curve. The Extremal Type I curve of Figure 1 was derived from the same data as the Weibull curve.

Joint Probability

It is often important to describe an event by more than one variable, such as both the duration and peak intensity, in which case the joint probability density must be evaluated. The probability that variables describing the event fall within specified ranges is determined from the joint probability density in a similar manner as with single variable probability density functions:

$$P(a_1 < x_1 < b_1, a_1 < x_2 < b_2, ..., a_n < x_n < b_n)$$

$$= \int_{a_1}^{b_1} \int_{a_2}^{b_2} \dots \int_{a_n}^{b_n} f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n$$
 (24)

when
$$f(x_1, x_2, ..., x_n) \ge 0$$

and

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n = 1$$
 (25)

A joint distribution function can be defined also:

$$F(x_1, x_2, ..., x_n) = \int_{-\infty}^{x_1} \int_{-\infty}^{x_2} \int_{-\infty}^{x_n} f(t_1, t_2, ..., t_n) dt_1 dt_2...dt_n$$
 (26)

The marginal probability density of variable $\mathbf{x_i}$ is determined by integrating the joint probability density function over the entire domain of all variables except $\mathbf{x_i}$:

$$f(x_i) = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(x_1, x_2, \dots, x_n) dx_1 \dots dx_{i-1} dx_{i+1} \dots dx_n$$
 (27)

An important feature of joint probabilities is that if the random variables involved are independent, then their joint distribution function is the product of their marginal distribution functions, such that:

$$F(x_1, x_2, ..., x_n) = F(x_1) F(x_2) ... F(x_n)$$
 (28)

Another important concept of joint probabilities is conditional probability density, defined in the case of two random variables as the conditional probability density of the first, \mathbf{x}_1 , given that the second takes on a specified value, \mathbf{x}_2 , or:

$$g_1(x_1 \mid x_2) = \frac{f(x_1, x_2)}{f(x_2)}$$
, if $f(x_2) = 0$ (29)

Conditional distribution functions, such as $F(x_1 \mid x_2)$, also can be defined, expressing the cumulative probability density in a manner analogous to single variable density functions. Conditional probability densities or distribution functions do not require independence for their definition.

Concepts Related to Evaluation of Risk

A traditional measure of risk of encountering an event of a specified intensity x, such as a critical flood elevation, wind velocity, or wave height, is the return period, RT(x). This is defined in practical terms as the average waiting period between exceedances of x. The return period for variables whose rate of occurrence is independent of their intensity (i.e., the number of occurrences per unit time is a Poisson process with a mean λ) is given by (Borgman and Resio 1982):

$$RT(x) = \frac{1}{\{\lambda[1 - F(x)]\}}$$
 (30)

The nonencounter probability, NE(x), is defined as the probability that, during a specific time interval L, the largest intensity encountered will be less than or equal to x. This can be expressed in terms of the

distribution function F(x) for the case of a Poisson process as (Borgman and Resio 1982):

$$NE(x) = e^{-\lambda L[1-F(x)]}$$
(31)

Expressed in terms of the return period:

$$NE(x) = e^{-L/RT(x)}$$
 (32)

This last relation demonstrates the danger of misinterpreting the return period as a frequency of occurrence for events of intensity x. When L = RT(x), then NE(x) = 0.37. In other words, there is a 63 percent probability of encountering an event of intensity x during the time interval L. The term "risk" is defined as the probability that an event of intensity x or greater will occur at least once in the time interval L, which is 1 - NE(x).

Another concept important in risk and optimization analyses is that of expectation, $E\{x\}$. This has actually already been defined as the mean of f(x):

$$E\{x\} = \mu = \int_{-\infty}^{\infty} xt(x) dx \qquad (33)$$

One useful feature of the expectation as a long-term average of the values of x is that the expectation of a function of x, g(x) can be defined by:

$$E[g(x)] = \int_{-\infty}^{\infty} g(x) f(x) dx \qquad (34)$$

Another feature with respect to Poisson processes worth noting regards the reference time period for risk criteria, such as estimation of the average annual value of some variable. Relation of the Poisson parameter λ to expectations of functions of the random variable x (the outcome of a Poisson process, where the number of occurrences per unit time is independent of the value of x) is easiest demonstrated by an example. Assume that in 1 year k

extreme events occur, where k is a Poisson variable. Intensities of extreme events are represented by significant wave heights, H_{si} (i = 1, 2, 3, ...k). Damage to a structure caused by each extreme event is assumed to be a function of H_{si} , $D(H_{si})$. Total damage in the year's time is:

$$\frac{D}{yr} = \sum_{i=1}^{k} D(H_{si})$$
 (35)

Since k and $H_{\mathbf{S}\mathbf{i}}$ are independent, then the expectation with respect to $H_{\mathbf{S}}$ is:

$$E\left(\frac{D}{yr}\right) = \sum_{i=1}^{k} E[D(H_{si})]$$
 (36)

Since H values are independent identically distributed random variables, they all have the same expectation, and:

$$E\left(\frac{D}{yr}\right) = E(k) E[D(H_g)]$$
 (37)

Taking the expectation of k to be the average number of extreme events per year (= the Poisson parameter, λ), the long-term average annual storm damage is:

$$E\left(\frac{D}{yr}\right) = \lambda E[D(H_g)] = \lambda \int_{-\infty}^{\infty} D(H_g) f(H_g) dH_g$$
 (38)

This relation is critical in optimization of first costs against estimates of long-term maintenance costs.

Regression by the Method of Least Squares

An important part of many research efforts is the estimation of distribution parameters from measured data by regression using the method of least squares. Assumed linear relationships between an independent variable x and a dependent variable y of the form:

$$y = \alpha + \beta x \tag{39}$$

can be tested against a set of x , y data and the differences, ϵ , between the estimated y and the predicted value measured. These differences can be due to measurement errors or inadequacies in the assumed relationship, such as neglect of other independent variables which also affect the value of y . The method of least squares allows the parameters α and β to be estimated by constants a and b such that resulting differences in the predicted versus measured y values are a minumum. Since these differences, called residuals, could be both positive and negative and therefore have a tendency to offset each other, the square of the differences is minimized instead. Many nonlinear relationships can be transformed into a linear form to take advantage of this technique.

The accuracy or reliability of least squares estimates of the true linear parameters α and β can be expressed in a number of ways. All possible true y values are assumed to be independently normally distributed with means $\alpha+\beta x$ and the common variance σ^2 . Measured values then can be written as:

$$y_{i} = \alpha + \beta x_{i} + \varepsilon_{i}$$
 (40)

where ϵ_i represents independent normally distributed random variables with zero means and a common variance σ^2 . This variance for "n" y values can be estimated in terms of the residuals as:

$$s_{e}^{2} = \frac{1}{n-2} \sum_{i=1}^{n} \left[y_{i} - (a + bx_{i}) \right]^{2}$$
 (41)

where $s_{\bf e}$ is the standard error of estimate. The standard error is in units of y and represents the limit within which approximately 68 percent of the absolute values of all errors will fall. Another quantitative measure of variance is the sum of the square residuals, or (n-2) $s_{\bf e}^2$.

The proportion of the variation of y values which can be attributed to the assumed relationship with x can be estimated as the ratio of the sum of squared residuals, $y - \hat{y}$, to the sum of squared deviations of y from the measured mean, y, subtracted from 1, the square root of which is known as the nonlinear correlation coefficient, r:

$$\mathbf{r} = \sqrt{1 - \frac{\Sigma(\mathbf{y} - \hat{\mathbf{y}})^2}{\Sigma(\mathbf{y} - \overline{\mathbf{y}})^2}}$$
 (42)

The above relation has the advantage over other correlation formulas that it is not restricted to linear relationships, although it is more tedious to compute.

Confidence that can be placed on predictions made with an equation developed by the least squares method can be estimated by various methods (Miller and Freund 1985, Isaacson and MacKensie 1981). The upper limit of confidence in estimates applied as design criteria always should be addressed by engineers as an integral part of the design process, particularly if predictions are extrapolated beyond the range of measured data. Techniques for estimating statistical confidence are not discussed here in detail since this project does not directly involve extrapolation. It should be noted, however, that obtaining a large sample is very important in improving statistical confidence. LeMehaute and Wang (1984 and 1985) have made special note of the sensitive effect on confidence of wave statistics attributable to the number of years of record and frequency of recordings. Neglect of statistical confidence inherent in formulation of structural design criteria can lead to inadequate safety and higher than anticipated maintenance costs for structures involved. The 20 years of hindcast wave data at 3-hr intervals available from the WIS program are valuable in this regard.

CHAPTER III: CHARACTERIZATION OF WAVE CONDITIONS

Basic Sinusoidal Concepts

An understanding of the basic theory and terminology of water wave mechanics is necessary for interpretation of hindcast wave information and any analytical application of this information. Water surface waves are most easily described as wave forms of sinusoidal shape. Certain key terms with reference to this simplified concept of water waves, as illustrated in Figure 2, include:

- 1. Wave height, H the vertical distance between a consecutive trough and crest
- 2. Wave length, L the horizontal distance between two consecutive crests (or troughs)
- 3. Wave period, T visualizing the wave form as travelling horizon-tally, the time for two consecutive crests (or troughs) to pass a fixed point, usually in seconds
- 4. Wave frequency, f nominally, the rate at which consecutive crests (or troughs) pass a fixed point (= 1/T), in hertz (cycles per second)

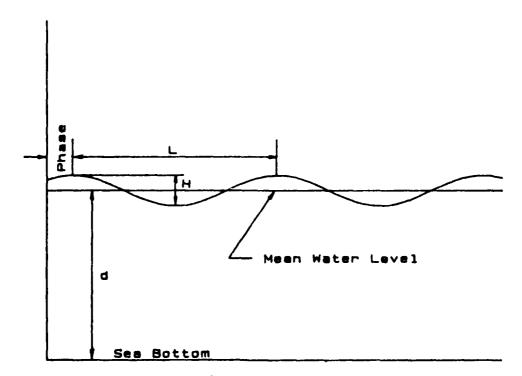


Figure 2. A sinusoidal wave

- 5. Radial frequency, ω the radial equivalent of frequency ($\omega = 2\pi/T$), also in hertz
 - 6. Wave number, k the radial equivalent of wave length (= $2\pi/L$)
- 7. Phase, ϕ the radial equivalent of the horizontal displacement, x', of a wave crest from the origin of the reference axis at time, t=0 (= $2\pi/x'$)

The basic equation which defines the wave profile in these terms is:

$$\eta(x, t) = \frac{H}{2} \cos (kx - \theta t + \phi)$$
 (43)

where $\eta(x, t)$ is the instanteous position of the water surface. Consideration of the sum of potential and kinetic energy inherent in a travelling wave of this form (per unit surface area) can be estimated by:

$$E = \frac{\rho g H^2}{8} \tag{44}$$

where ρ is the mass density of the seawater. This total energy is notably a function only of the wave height squared (Dean and Dalrymple 1984).

A consideration of surface, bottom, and transverse boundary conditions, with simplifications which eliminate all but first-order differential terms, yields the mathematical equation, known as the dispersion relation, which predicts effects of depth on wave length:

$$\omega^2 = gk \tanh (kd) \tag{45}$$

where g is the acceleration due to gravity and tanh is the hyperbolic tangent. A feature of sinusoidal waves which is consistent with this relation is that deepwater wave length, $L_0 = (g/2\pi) T^2 = 5.12T^2$ ft or $1.56T^2$ m. The speed at which a wave crest travels, the phase velocity, C, in deep water = $L_0/T = 5.12T$ ft/sec or 1.56T m/sec. The change that occurs in shallower water is that wavelength shortens and phase velocity, C = L/T, increases. The wave height also is affected, first slightly decreasing, then increasing as the water grows more shallow. The overall tendency of water waves to

change form as depths decrease is known as shoaling. The change in wave height due to shoaling is governed by:

$$K_{s} = \frac{H}{H_{o}} = \left(\frac{C_{o}}{2C_{g}}\right)^{1/2} \tag{46}$$

where H and H are shoaled and deepwater wave heights and K is the shoaling coefficient. The variable $C_{\rm g}$ is the shoaled group velocity, the speed at which groups of waves travel which is also the speed at which wave energy approaches shore:

$$C_g = \frac{C}{2} \left[1 + \frac{2kd}{\sinh(2kd)} \right]$$
 (47)

where C is the shoaled phase velocity (= L/T) and sinh is the hyperbolic sin function (Dean and Dalrymple 1984).

The wave form becomes steeper in decreasing depths, ultimately reaching an unstable state when breaking occurs. The point at which breaking actually occurs is not fully understood at this time, but, based on the theory of solitary waves, generally occurs at the point where the wave height, H = 0.78d. Some field data tend to show that most locally wind-generated waves (i.e. "seas") break in deeper water, with breaking heights on the order of 0.6d to 0.7d. Very long waves not locally generated (i.e. "swell") may not break until they are in very shallow water, however, since they may form surging breakers analogous to hydraulic phenomena known as "bores" or "hydraulic jumps."

The discussion above is meant to point out that there are practical limits to wave heights at most coastal sites due to breaking, but that these limits are as yet difficult to reliably define in practice. Furthermore, simplifications inherent in first-order sinusoidal theory are not sufficiently accurate for engineering purposes in many shallow-water situations and predictions made with a higher order wave theory must be applied.

Shoaling occurs only as a function of depth, but refraction also affects the wave form as a function of wave direction with respect to depth contours of the sea bottom. Refraction of water waves is analogous to refraction in classical physics of a ray of light passing through a pane of glass at an

angle. The most frequently observed effect of water wave refraction is for waves approaching the coast at an angle to bend around as their crests tend to become parallel to the shoreline in shallow water. Snell's Law is usually applied to describe the change in angle of water waves by refraction in much the same way as it is in optics, commonly stated as:

$$\frac{\sin \theta}{C} = \frac{\sin \theta_{0}}{C_{0}} \tag{48}$$

where C and C are the refracted and deepwater phase velocities (= L/T and L /T) and θ and θ are the refracted and deepwater angles of wave crests with the bottom contours. Snell's Law assumes straight and parallel contours between deep water and the depth at which the above relation is applied. The relation can be applied in increments of incident versus refracted angles and thus applied to gently curving contours. Refraction usually (except in cases of convergence at convex contours) causes a reduction in wave height, which is superimposed on the effect of shoaling, according to the ratio:

$$K_{r} = \frac{H}{H_{o}} = \left(\frac{\cos \theta_{o}}{\cos \theta}\right)^{1/2} \tag{49}$$

where H and H are the refracted and deepwater wave heights and K_r is the refraction coefficient (Dean and Dalrymple 1984).

Wave diffraction describes the effect which a partial barrier has on wave heights beyond the barrier. It is the process which allows wave energy to leak sideways behind an obstruction or laterally from an area of high energy to an adjacent area of lower energy. The head of a breakwater, for example, will cause waves to diffract behind the breakwater into its geometric shadow, even though it may prevent any other form of wave transmission. Larger scale landforms and submerged formations can cause a degree of wave diffraction. Precise predictions of the effects of diffraction are more complicated than for shoaling and refraction, but the combined effects of these three forms of wave transformation are important in explaining observed behavior of water waves in many practical situations. The complexity of

diffraction often requires the use of physical scale models to ensure with confidence satisfactory performance of protective structures such as break-waters enclosing a port or harbor area.

Irregular Waves

The fact that real ocean waves typically appear chaotic with little regular form was mentioned previously. An explanation of this reality is that wave groups from many different sources with different heights, periods, phases, and directions are interacting in the small area we observe with the resulting superpositions appearing as chaos. Figure 3 illustrates a

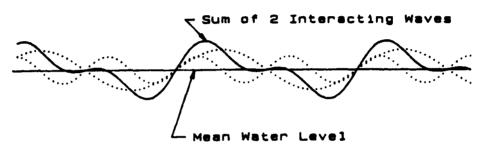


Figure 3. Interaction of sinusoidal waves

hypothetical point in time when two sinusoidal wave groups interact, one with 50 percent greater height and period and a $\pi/4$ phase difference. The waves would appear criss-crossed when viewed from above if their directions were not parallel.

Actually, winds that create the waves generate a range of heights and periods. Since phase velocity varies with period, longer period waves travel faster and soon leave shorter period waves behind. Swell, as previously defined, refers to waves which have completely left the area in which they were generated. These waves typically have periods greater than about 9 or 10 sec, but a clear distinction does not exist. Waves which are still within the influence of the generating wind system are called "seas" and typically are dominated by shorter period waves (less than 9 sec).

The distribution of individual wave heights in a stationary sea state has been found in most cases to follow a Rayleigh distribution, as discussed in the previous paragraphs on statistical concepts. Stationarity technically is the condition during which all moments (including the mean and variance)

are time invariant (Bendat and Piersol 1971). A small sample thus can be analyzed and taken to represent the entire period during which conditions remain stationary. Waves at sea are assumed by most investigators to be weakly stationary for periods of about 3 hr, occasionally for as much as 6 hr, but seldom longer. This is more of a tradition related to the practicalities of collecting wave data than a precisely defined interval. The parameters derived from an instanteous measurement (such as the case of synoptic hindcasting) or from a 20-min recording of the water surface elevations are therefore typically taken to represent a much longer period during which conditions do not change. This, of course, is not really true, but as long as the changes are not drastic and are generally within the confidence limits of the statistical parameters of interest, this practice is acceptable.

Wave periods do not lend themselves as readily as do wave heights to representation by a standard statistical distribution such as the Raleigh distribution. Bretschneider (1959), however, found that the distribution of squared wave periods, T^2 , for seas followed a Rayleigh distribution. Other investigators have applied a variety of standard distributions, and specialized empirical distributions also have been developed.

The practice of coastal engineers in the last 10 years has largely shifted from considerations of wave period exclusively in the time domain to frequency domain considerations. Decomposition of a time series of water surface elevations into a set of incremental sinusoids, each represented by an amplitude (= H/2) and a frequency (= 1/T), can be accomplished by transformation of the time series into its equivalent Fourier series. Wave conditions thus can be represented by the distribution of wave energy (proportional to amplitude squared per Equation 44) as a function of frequency, or a wave spectrum.

Figure 4 illustrates a wave spectrum with two "peaks," one representing swell-type waves and the other representing coexistent seas. The inverse frequency of the dominant peak is in practice usually taken as the peak period, which is generally assumed as the most probable period in the sea state. This is a "one-dimensional" spectrum which does not account for the direction of wave energy propagation. More complex procedures have been developed to express the distribution of wave energy as a function of both frequency and direction. The most common practice is to treat the directional spread of wave energy to be independent of the distribution of energy by frequency. This

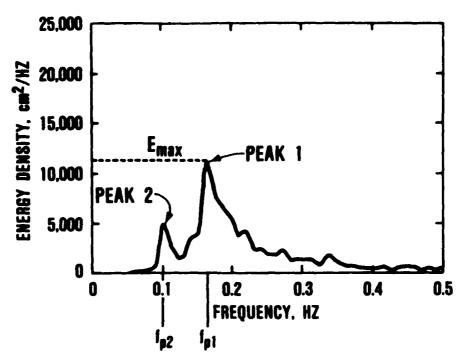


Figure 4. An example of a double-peaked energy density spectrum

allows application of a spreading function $\theta(\theta)$ which, when multiplied by the one-dimensional spectrum S(f), yields the two-dimensional spectrum $S(f,\theta)$:

$$S(f, \theta) = S(f) \theta(\theta)$$
 (50)

The form of a spectrum is quite sensitive to the analytical procedures applied, particularly "smoothing" performed to improve statistical confidence at the cost of resolution. Most spectral analysis procedures actually deal with discrete frequencies (= $2\pi/T$ of the individual sinusoids) which, when averaged over equal intervals, yield a smoother looking plot with more narrow confidence bands. A jagged looking spectrum will have wider confidence limits than a smoothed spectrum computed from the same data.

Integration of a wave spectrum which has been computed as energy per frequency band, $E/\Delta f$ (e.g. m^2/Hz), versus frequency yields the total energy of the sea state. This relates directly to actual variance of the water surface elevations such that:

$$\sigma_{ws}^2 = \int S(f) df$$
 (51)

where σ_{ws}^2 is the variance of the water surface elevations and S(f) is the computed energy density spectrum. Spectra in this form are often taken as continuous functions since it is reasonable to expect wave energy to be generated in continuous frequencies.

A parameter in units of wave height which has been used to represent the range of wave heights in a sea state is the zero moment wave height, $H_{mo} = 4\sigma_{ws}$. The "zero moment" title comes from integration of $f^{n}S(f)$ with respect to f where n, the power of f in the integral, is zero as with Equation 51. This wave height has been found to be very close to the significant wave height, H_{s} , of Rayleigh distributed seas in deep water. H_{s} typically departs from H_{mo} in shallow water (Thompson and Vincent 1983). The zero moment wave heights corresponding to two interacting wave groups of double-peaked energy density spectra, as illustrated in Figure 4, can be estimated by splitting the spectrum between peaks and integrating each side separately. There is no widely accepted way to estimate the parameters of multiple wave groups from their combined spectrum, but this method gives an indication of their relative intensity as potential structural design criteria.

A number of parameterized spectra have been developed in the effort to relate wave conditions to winds and geographical factors which constrain generation of waves at sea. These parametric spectral forms nearly all apply to waves in the generation phase, i.e. seas, not swell. The four most important factors in wave generation are wind velocity (and resultant stress) over water, duration of that velocity, fetch (distance over water which the wind blows), and water depth. Depth limitations on wave spectra are the most recent effects to be reliably defined in combination with other primary constraints. Other factors which also can be significant are preexisting waves (wave-wave interaction) and the presence of strong currents (wave-current interaction). Waves generated by winds of a given velocity in water of a given depth thus are either duration limited, fetch limited, or fully developed and may be affected by waves coming into the generation area from a distant source and strong currents. Virtually all parametric spectral shapes have the "tail" of the spectrum, the portion to the right of the peak,

proportional to f^{-5} , following the work of Phillips (1977). An advanced form, as an example, is the TMA spectrum (Hughes 1984), which includes the depth limitation:

$$S(f, d) = \alpha g^{2} f^{-5} (2\pi)^{-4} \phi (2\pi f, d) e^{-5/4 (f/f_{p})^{-4}} \exp -(f/f_{p}^{-1})^{2} / 2\sigma_{\star}^{2}$$
(52)

where $\phi(2\pi f, d)$ is a function of depth (d), k (the wave number, $2\pi/L$), and ω (the radial frequency, $2\pi/T$) allowing portions of the spectrum to be transformed by linear wave theory. The term α is the Phillips equilibrium constant, which has recently been taken to be a function of depth, wind speed, and peak frequency, f_p . The γ term is the "shape parameter" which is a function of wind speed and fetch. The σ_{\star} term is an empirical factor affecting shape of the spectrum on either side of the peak. This form applies to fully developed or "saturated" seas in decreasing depths. Figure 5 illustrates the effect of changing depth on TMA spectral shape. The deepwater predecessor of the TMA spectrum, the JONSWAP spectrum, now is widely used to predict both fetch and duration limited wave growth in deep water (Vincent 1984).

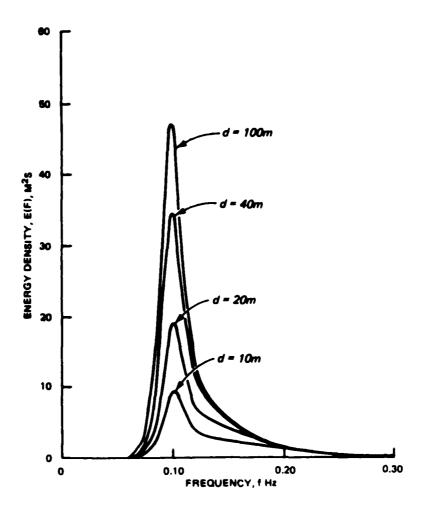


Figure 5. The TMA spectrum

CHAPTER IV: WAVE INFORMATION STUDIES HINDCAST DATABASE

General Background of Phases I and II

The WIS program of the US Army Engineer Waterways Experiment Station began in 1976 with the goal of providing a long-term (20-year) hindcast of wave information for use in development of design criteria for coastal projects. The term "hindcast" refers to the technique of simulating historical wind and wave generation from pressure data available from surface weather charts. The basic raw data for hindcasting thus are instanteous pressure recordings which meteorologists have applied to produce pressure fields delineated by isobars and other notation common to surface weather charts. These "highs," "lows," "fronts," "troughs," and "ridges" are then applied to simulate the effect of corresponding wind fields on the surface of the ocean.

The WIS program first transcribed into digital form pressures from surface weather charts from 1956-1975 for the North Atlantic, Gult of Mexico, and North Pacific, with as much checking for accuracy and consistency as the basic data allowed (Corson, Resio, and Vincent 1980). This information was available at 6-hr intervals. Winds which would have existed with each consecutive distribution of pressures next were simulated by a series of numerical models assuming quasigeostrophic flows and a planetary boundary layer which yielded surface level (19.5-m elevation) wind fields. These wind fields were in turn adjusted with observations of actual wind velocities, wherever possible (Resio, Vincent, and Corson 1982).

Given the database of surface level winds created by the steps above, basin geometry and grid were defined for numerical simulation of deepwater wave generation. Figures A-1 and A-2 illustrate deepwater (Phase I) grids for the North Atlantic and North Pacific Oceans. Execution of a deepwater numerical model of wave generation, which took into account fetch, duration, directional spreading effects, and wave-wave interaction, produced a database of two-dimensional spectra and related parameters at intersections of the grid lines. Detailed wave information was retained only at intersections marked with dots and published in written form (Corson et al. 1981 and Ragsdale 1983) tor the numbered sites.

Phase II of the WIS program performed simulations at 3-hr intervals of wave generation in a manner similar to Phase I (deep water), but at a finer

scale and in transitional depths of the continental shelf. Figures A-3 and A-4 show the Atlantic and Pacific Phase II grids and stations where wave information has been published (Corson et al. 1982 and Ragsdale 1983). In addition to Phase I factors, Phase II simulations took into account the sheltering effect of large-scale land masses, refraction, and shoaling. The Phase I wave information served as a boundary condition at the seaward limit of the Phase II grid.

Neither Phase I nor Phase II distinguished seas and swell, but rather dealt with individual discrete frequency bands over the entire two-dimensional spectrum at any point. Phase III decomposed this spectrum into seas and swell, treating seas as two-dimensional spectra and swell as monochromatic, unidirectional wave groups. The definition of swell as waves which have travelled beyond the area in which they were generated was applied. This approach economized computations by taking advantage of the fact that swell typically has its energy highly concentrated in a narrow band of frequencies, which is close to a monochromatic condition. Wave parameters computed and recorded in the Phase III database included zero moment wave height, peak period, and dominant direction of propagation. Monochromatic equivalents were recorded in the case of swell and combined wave heights were recorded as:

$$H_{combined} = \sqrt{H_{sea}^2 + H_{swell}^2}$$
 (53)

Period and direction recorded in the "combined" category corresponded to the peak period and dominant direction of either the sea or swell, whichever had the higher zero moment wave height (Brooks and Corson 1984). The Phase III approach is most valid for coasts with straight and parallel contours and is less precise in more complex bathymetry.

Phase III Shallow-Water Wave Information

Phase III efforts of the WIS program were directed at providing wave information suitable as design criteria for a great many coastal endeavors in a depth of 10 m at 10-mile (16.1-km) intervals along the Atlantic (Jensen 1983a) and Pacific (Ragsdale 1983) coasts of the continental United States. This task dealt with transformation of wave conditions from Phase II stations to

166 Atlantic and 134 Pacific Phase III stations. Figure A-5 illustrates a section of the Atlantic Phase III stationing system and adjacent Phase II stations. The magnitude of data processing requirements and complexity of the coast at this finer scale led to procedures for estimating wave conditions in shallow water (10 m depth) described briefly below.

A spectral (frequency domain) approach to wave transformation was sought to reduce computational time required to simulate wave transformation in the time domain. A parameterized spectrum was necessary for this, but one as complex as the TMA spectrum, or the most refined spectral forms available at the time of the Phase III procedure formulation, would have provided an unmanageable computational burden. The one-dimensional parameterized spectrum chosen for Phase III simulations had the following form:

$$S(f) = \alpha_g^2 f^{-5} (2\pi)^{-4}$$
 for $f \ge f_p$ (54)

$$S(f) = \alpha g^2 f_p^{-5} (2\pi)^{-4} \exp \left[1 - \left(\frac{f}{f_p} \right)^{-4} \right] \quad \text{for } f < f_p$$
 (55)

which applied the well-accepted f^{-5} right-hand tail, but limited free parameter determination to only two variables, α and $f_{\rm p}$ (Kitaigordskii 1962).

A spreading function, assumed to be independent of the one-dimensional spectral form, was defined as:

$$\Theta(\theta) = \frac{8}{3\pi} \cos^4 (\theta - \theta') \tag{56}$$

where e' is the predominant direction of propagation. Thus, the twodimensional form was:

$$S(f, \theta) = S(f)\theta(\theta) \tag{57}$$

Within each 10-mile (16.1-km) interval defined as Phase III stations along the coast, bottom contours were assumed to be straight and parallel. A specific orientation was assigned to each interval such that departure of this assumption from the true situation was minimized. The processes of refraction

and shoaling, as defined by Snell's Law and sinusoidal theory, were applied to increments of trequency and direction of the directional distribution defined by $S(f, \theta)$. Wave energy propagating seaward was ignored.

The geometric relationship between a Phase III station and adjacent Phase II stations from which the model derived its input was the most important consideration in addressing sheltering in Phase III. Basically, the geometric shadow of a landform to wave energy from a specific direction was considered as absolute, i.e., no energy was propagated into the shadow area. This is a gross simplification, but it made the simulation of sheltering effects practical for Phase III. Discrete combinations of frequency and direction were considered incrementally with respect to sheltering, as they were with refraction and shoaling.

The problem of wave-wave interaction and the losses it can cause, evidenced by white caps and other signs of turbulent energy dissipation, was addressed by definition of another spectral form for shallow water. Principles of similarity were applied to derive a form consistent with Phase I and II deepwater considerations, which predicted the spectrum in shallow water:

$$S(f) = \alpha g h(8\pi)^{-2} f^{-3}$$
 for $f \ge f_p$ (58)

This relation is consistent with the visualization that energy losses due to wave-wave interaction tend to occur at high frequencies, while energy at lower frequencies is conserved. A further application of equilibrium principles allowed derivation of an integrated form of this equation which describes the dependency of sea wave heights on depth:

$$(H_{\text{seas}})_{\text{max}} = \frac{(\alpha \text{gd})^{1/2}}{\pi f_{\text{c}}}$$
 (59)

where $f_c = 0.9f_p$ is a energy cutoff frequency (lower integration limit) and (H seas max treated differently for swell, however, in the manner of estimating breaker heights for monochromatic waves. A breaking coefficient of 0.6 was applied, which is consistent with recent measurements of breaking waves by the WES (Jensen, Robert E., verbal communication, February 1986):

Extensive comparisons have been made between the limited measured wave data available and WIS wave information, generally with acceptable results (Corson and Resio 1981). The reduction of measurements made by wave gages also involves compounded assumptions, and discrepancies between wave information based on gage data and Phase III wave information could not always be resolved. More accurate techniques are available for site-specific simulation of the transformation of waves into shallow water. These methods unfortunately were too complex to apply systematically on the scale of the WIS Phase III endeavor, though improvements are under consideration. The presently available end product of Phase III is, however, an excellent tool for coastal engineers to use in the planning and preliminary design stages of coastal projects for development of design criteria. More complex and expensive numerical simulations and physical scale models can be performed in the detailed design phase after the economic feasibility and financeability of the project has been ensured. Even in the final stage, some basis of experiment design and cross-check on other sources of wave information is necessary. The 20-year period of record for the WIS database can rarely be exceeded by other reliable sources. The WIS wave information provides, therefore, a vast improvement to the confidence of each design effort to which it is applied.

CHAPTER V: LITERATURE REVIEW OF STORM DURATION STUDIES

Recent Literature on the Duration of Sea States

Table 1 presents mean durations for various weather types in the British Isles which were excerpted from Barry and Perry (1973). The weather type

Table 1
Mean Durations of Weather Types in the British Isles

Mean Duration (days)				
January		July		
1910-1930	1948-1968	1910-1930	1948-1968	
4.1	1.7	2.6	2.7	
1.5	1.8	2.0	1.8	
1.9	2.6	2.0	1.8	
2.0	1.8	1.7	1.3	
1.4	1.3	1.7	1.9	
2.2	1.9	2.5	2.2	
	1910-1930 4.1 1.5 1.9 2.0 1.4	January 1910-1930 1948-1968 4.1 1.7 1.5 1.8 1.9 2.6 2.0 1.8 1.4 1.3	January 1910-1930 1948-1968 1910-1930 4.1 1.7 2.6 1.5 1.8 2.0 1.9 2.6 2.0 2.0 1.8 1.7 1.4 1.3 1.7	

identified as "cyclonic" is assumed to meet the standard definition of winds circulating around a low pressure area (Lester 1973), corresponding to the extratropical cyclonic events which are simulated in the WIS program. This type of weather is noted to have a mean duration of 1 to 2 days in Great Britain, with some seasonal variation. Statistics of this type would surely vary from region to region, but the order of magnitude in hours, say less than 100 but more than 10, can serve in this investigation of storm characteristics as a rough first measure of a reasonable mean duration. The untrained intuition of any regular viewer of television weather reports would likely agree with this typical range.

Surprisingly little material was available in the coastal engineering and oceanographic literature which dealt directly with the duration of extreme events at sea or of extreme wave conditions. Occasional references were made to a 3-hr period of wave height stationarity assumed for practical purposes in measurement programs (e.g., Agerschou et al. 1983 and Massie 1976). The interval between samples of wave measurements is commonly set at 3 hr.

Publications of WIS wave information (Corson et al. 1981 and 1982 and Jensen 1983a) tabulated durations of significant wave heights above selected thresholds, but did not discuss trends or other implications inherent in this information.

North Sea Investigations of Houmb and Vik

The most rigorous work to date has been a series of studies by two Norwegian investigators (Houmb 1971, Houmb and Vik 1975, Vik and Houmb 1976, and Houmb and Vik 1977). Other authors have reviewed this work (e.g. Battjes 1977, PIANC 1979, and Bruun 1985), but no significant advances seem to have been made regarding the characterization of extreme event durations following Houmb and Vik (1977). Their work on the duration of sea states culminated in the findings of the last reference, which will be reviewed in detail in the following paragraphs.

Houmb and Vik (1977) considered both the duration of extreme events, specified as the time during which the significant wave height exceeded a given threshold, and the duration of "calms" between these extremes. The basis of their investigations was wave recordings made at five North Sea sites where depths varied from 80 to 250 m. Three sites involved time series measurements made for 20 min every 3 hr. A fourth site involved 10-min time series measured every 4 hr. The sequences of these measurements were not continuous and varied in total period of record from 3 to 31 months. The fifth site provided observations from a rescue vessel every 3 hr from 1959 to 1974 during October through March only. These observations classified predominant wave heights into classes of 0.5 m.

A theoretical approach toward prediction of variation of storm durations was first proposed by Houmb and Vik (1977) which took the frequency, or marginal probability density, of threshold up-crossings (i.e. $H_s^i = dH_g/dt$ was positive) as:

$$f(H_t) = H_s'f(H_t, H_s') dH_s'$$
(61)

where H_t is the specified threshold and $f(H_t, H_s')$ is the joint probability density of H_s and its time derivation, H_s' . The average duration of extreme events, $t(H_t)$, $(H_s > H_t)$ was derived to be:

$$t(H_{t}) = \frac{L[1 - F(H_{t})]}{f(H_{t})L} = \frac{[1 - F(H_{t})]}{f(H_{t})}$$
(62)

where L is the period of interest (say 50 years) and $F(H_t)$ is the cumulative distribution of H_s evaluated at H_t , or the probability that H_s is equal to or less than H_t . The quantity $[1-F(H_t)]$ is the probability that H_s is greater than H_t . The average number of up-crossings, i.e. the average number of extreme events, during the period L was taken to be $f(H_t)L$, where $f(H_t)$ is the probability density of H_s at H_t given above.

The rate at which H_s changes (from one stationary period to the next) was assumed to be a Poisson process, i.e. H_s^{\dagger} was assumed to be independent of H_s . The joint probability density function $f(H_s, H_s^{\dagger})$ could then be evaluated as:

$$f(H_s, H_s') = f(H_s)f(H_s')$$
(63)

The marginal probability density function $f(H_S)$ was assumed to follow a Weibull distribution whose corresponding distribution function had the form:

$$f(H_s) = 1 - \exp\left(-\frac{H_s - H_o}{H_c - H_o}\right)^T$$
 (64)

and

(65666333) (6373333)

$$f(H_s) = \frac{T(H_s H_o)^{T-1}}{(H_c - H_o)^{T}} \exp\left(-\frac{H_s - H_o}{H_c - H_o}\right)^{T}$$
 (65)

where H_{c} , H_{o} , and T are parameters of the distribution.

The function $f(H_s^i)$ was assumed to be normally distributed with zero mean for positive values of H_s^i (increasing H_s). The data seemed to support this assumption. This gave $f(H_s^i)$ as:

$$f(H_s^{\dagger}) = \frac{1}{2\pi\sigma_h} \exp\left(\frac{-H_s^{\dagger 2}}{2\sigma_h^2}\right)$$
 (66)

where σ_h is the standard deviation of H's which was evaluated from the data. The differences between σ_h values computed for increasing and decreasing H s were found to be negligible. Furthermore σ_h was not noted to follow a seasonal pattern. This application of the above normal distribution with zero mean gave the advantage of requiring only one parameter, σ_h , to be determined empirically, in addition to those (H , H , and T) for F(H s). The resulting function for the mean duration $t(H_+)$ reduced to:

$$t(H_{t}) = \frac{2\pi (H_{c} - H_{o})^{T}}{T\sigma_{h} (H_{t} - H_{o})^{T-1}}$$
(67)

The cumulative distribution of measured durations was found to be well represented by a Weibull distribution of the form:

$$F(t) = 1 - \exp \left[-\left(\frac{t}{t_c}\right)^{\alpha} \right]$$
 (68)

where α is the shape parameter and t_c is the scale parameter. Average durations estimated by the $t(H_t)$ function derived above also compared well with means computed from the set of measured durations. Houmb and Vik (1977) gave examples of how this formulation could be applied in the conduct of offshore oil explorations, as in prediction of duration of operation down time caused by extreme wave conditions.

The formulation of Houmb and Vik (1977) was well defended in terms of conceptual limits of parameters such as H_s^{\prime} and σ_h . They tested their hypotheses as well as possible with their limited data set, but urged in their conclusions that further investigations be pursued with more comprehensive wave intormation.

CHAPTER VI: EXTREME EVENT IDENTIFICATION

Choice of Sites

Each Phase III site includes 58,440 records of wave information 3 hr apart from 0000 (midnight) January 1, 1956, to 2400 (midnight) December 31, 1975 (20 years). Four sites were originally chosen for analysis, two on the Atlantic coast and two on the Pacific coast. A third Atlantic site was later chosen when it was discovered the first two had very similar distributions of significant wave heights. The five sites ultimately investigated are listed in Table 2. They were intended to represent a wide geographical spread in

Table 2
WIS Phase III Stations Investigated

Station	Site	Latitude	Longitude
A3061	Atlantic City, New Jersey	39.34° N	74.47° W
A3083	Nagshead, North Carolina	35.94° N	75.61° W
A3142	Daytona Beach, Florida	29.20° N	81.00° W
P3036	Newport, Oregon	43.63° N	124.08° W
P3105	Half-Moon Bay, California	37.45° N	122.45° W

hopes that analysis would reveal any important universal traits or significant geographical differences. Figure 6 shows their relative location along the US coasts. Statistics published by the WIS program (Jensen 1983a, b) for the Atlantic sites are presented in Appendix A. Wave height frequency tables (not yet published by the WIS program) for the two Pacific sites also are presented in Appendix A.

Basic Treatment of WIS Phase III Wave Information

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Table Al illustrates format and unit conventions of the WIS Phase III database. Dates are given as year/month/day and times referenced to the 24-hr clock (i.e., military time). Wave heights, i.e. the zero moment wave heights derived for each 3-hr time step, are reported in centimetres. Wave periods,

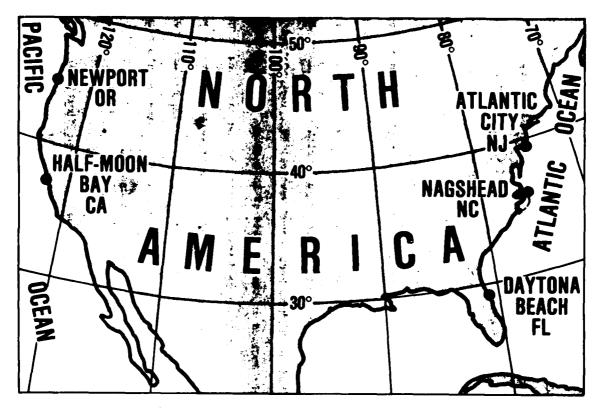


Figure 6. Geographic relation of sites investigated

i.e. the peak periods of the hindcast spectra, are reported to the nearest second. Direction or azimuth is reported in degrees relative to the shoreline, such that 90 deg is a wave direction travelling straight into the straight and parallel contours assumed for each 10-mile (10.1-km) shoreline increment. Combined statistics presented in Table Al include the geometric average wave height (Equation 53) and the peak period and predominant direction of either seas or swell, whichever had the highest zero moment wave height. Combined statistics were applied in analyses of this study, though they were not actually a part of stored wave information and had to be computed. Mean and maximum duration of exceedance of selected wave heights were reported for the Atlantic sites by Jensen (1983b) and are included in Appendix B. A comparison of those statistics with results from this investigation is made later in this report.

The Problem of Extreme Event Identification

The work of Houmb and Vik (1977) on duration of sea states was apparently performed exclusively with significant wave heights crossing an

arbitrary threshold. This implies that the significant (or zero moment) wave height is the most appropriate measure of the extreme events' intensities for applications of duration statistics. Other parameters can be conceived, however, which might be better representatives of the overall intensity or extreme nature of a storm. The most obvious alternate parameter would be peak period, to which refraction, shoaling, and wave breaking are all quite sensitive. Wave length might be another, although wave length at any depth is a function of period. Wave steepness, H/L, is commonly associated with breaker characteristics and forces on coastal structures. If the ratio of zero moment wave height to deepwater wavelength corresponding to the peak period is used, representative wave steepness becomes $2\pi H/gT^2$. The 2π factor is commonly dropped as a part of this dimensionless steepness parameter in favor of H/gT^2 .

Wave severity, H^2L , has recently become of interest as a factor closely related to stability of rubble-mound structures (Graveson et al. 1980 and Ahrens 1984). Wave severity can be thought of as the ratio of wave height cubed (the traditional wave parameter for evaluation of rubble-mound stability) to wave steepness, H/L. Again, significant or zero moment wave height and deepwater wave length corresponding to the peak period of the spectrum are used for convenience, yielding $\text{H}^2\text{L} = 2\pi\text{H}^2/\text{gT}^2$. It should be noted that the four parameters discussed so far vary the relative influence of wave height and period in the following order: H, T, H/T^2 , and H^2/T^2 . These parameters also could be used to define extreme event duration as the time during which consecutive parameter values exceed a specified threshold value.

A fifth parameter which might be important with respect to duration of extreme wave conditions is predominant wave direction. This certainly would be true for sites naturally protected in all but one narrow sector. WIS Phase III data did not include any such sites, however, assuming an open coast with sheltering only from major landforms.

Figure 7 illustrates the time series for wave heights during October 1956 at Nagshead, North Carolina. This particular time span was chosen for presentation because it included rapid changes in wave conditions, especially on October 27 and 28, 1956, as indicated by sharp spikes near the end of the wave height time series plot of Figure 7. Table Al includes Phase III wave information recorded for these 2 days.

Figure A6 shows the time series of peak wave period during this same

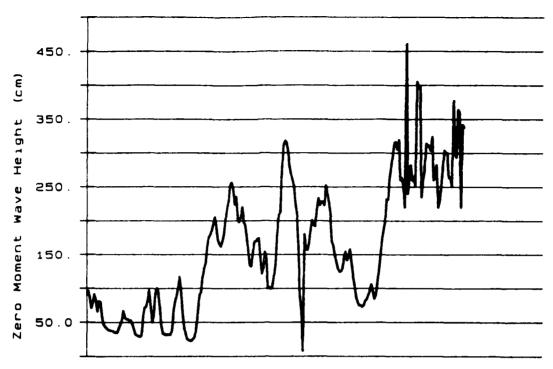


Figure 7. Wave height time series: Nagshead, North Carolina, October 1956

month for Nagshead, North Carolina. The wave period can be seen to vary somewhat out of phase with wave height and to have a tendency to remain constant for significant time spans and then change abruptly. A plot of H/gT for the same period at Nagshead (Figure A7) appears more like the wave period time series than the wave height time series, also tending to vary slowly for significant time spans and change abruptly (the influence of T2 in the denominator). A plot of wave severity, H²L (Figure A8), for the same time period dramatically delineates extremes of the wave height time series. When plotted in Figure A9 as $\left(\mathrm{H}^{2}\mathrm{L}\right)^{1/3}$, however, wave severity very closely resembles the wave height time series plot. Wave severity in this form has the same units as wave height and includes the influence of H2 to balance the influence of in the denominator. The plot of direction (Figure Al0) does not indicate direct relation to the wave height plot and is much more erratic, even in nonextreme periods. Direction can be considered to be practically independent of the intensity of wave conditions since it is controlled almost exclusively by geometric factors.

The convention of previous investigators (Houmb and Vik 1977) to rely solely on variation in zero moment wave height for definition of extreme event

durations was maintained in this study. This parameter is most easy to visualize and has a long tradition as the critical measure of intensity of extreme events at sea. Variations of $\mbox{H}^2\mbox{L}$ and $(\mbox{H}^2\mbox{L})^{1/3}$ show promise, but the large units of $\mbox{H}^2\mbox{L}$ make results of computations rather abstract and the variation of $(\mbox{H}^2\mbox{L})^{1/3}$ seemed quite close to that of \mbox{H} . Relationships of individual extreme event durations (measured by variation of \mbox{H}) to peak conditions measured by all parameters discussed above were investigated, however, and the results of that analysis are reported later in this report.

An investigation of actual weather conditions on the Atlantic coast in the time frame surrounding October 27-28, 1956, was conducted to better understand what events were actually driving the numerical simulations to produce irregularities in the time series of Figure 7. First, Phase II data input to the Phase III numerical wave transformation were inspected. Table A2 presents Phase 11 information at Station A2037, at 36.06° N latitude and 74.92° W longitude, approximately 33 nautical miles (61 km) east-southeast of Nagshead in about 240 ft (73 m) or water. The intermittent appearance and disappearance of swell can be seen to follow a similar pattern in the Phase III site of interest (Station A3083) and the Phase II site directly offshore (Station A2037). Wave heights in deeper water are higher, lacking the depth limitations inherent in Phase III simulations. Wave periods of both sites are identical, unaffected by the wave transformation processes simulated in Phase III. The direction convention in Phase II is different, indicating the direction from which waves are travelling toward the center of the compass rose. Phase II data do not include anything significantly revealing about the irregularities of interest, basically showing the same patterns in this case.

The nearest Phase I site offshore of Nagshead was Station A1005 at 35.4° N latitude and 72.3° W longitude, located in deep water approximately 163 nautical miles (302 km) east-southeast of Nagshead, North Carolina. Table A3 shows Phase I information recorded for October 26-28, 1956. There is only one record which included swell; that record did not dominate the combined wave height, which appears to be steadily decreasing at that time. It is important to recognize that a significant travel time would be involved between this Phase I site and Stations A2037 or A3083 (approximately 8 and 10 hr, respectively, for waves of 11-sec period), so the conditions at a given date and time should be "out of phase" by three to four records.

An inspection of surface weather charts during the later part of October

1956 for North America and the north Atlantic Ocean was made to identify synoptic weather systems which may have dominated Phase I and Phase II information. An explanation was sought for the sudden appearance and disappearance of swell in the data, as well as an explanation of differences between Phase I and Phase II wave information. Figure 8 illustrates recorded weather patterns of October 26-28, 1956, showing the presence of a generally stationary, weakly defined, low pressure system of fluctuating intensity offshore of Cape Hatteras. This location is close to Station Al005; thus, the basic definition of swell as waves which have left their area of generation could explain the lack of swell in Phase I data. The wave field at this point would have been under the influence of cyclonic winds of the low pressure system and thus only seas would have existed, as defined by WIS conventions. The relative position of Stations A2037 and A3083 in combination with the fluctuating intensity of the low pressure system appears to have caused swell either to come from too tar south to affect Nagshead or to exist only as seas, except for the spikes of Figure 7. This set of circumstances is probably exceptional, but an understanding of the real weather patterns driving numerical simulations of the WIS program in this instance may help explain trends of duration revealed by further analysis of WIS data.

Analytical Procedure and Results

A FORTRAN computer program was written which read the 58,440 records stored for each Phase III site and maintained a record of the number of consecutive records, each of which had a combined wave height above a specified wave height threshold, HI. Subsequent us of the term "extreme event" refers to events defined in this manner. The number of extreme events was counted and statistics including the maximum, minimum, mean, and standard deviation durations were computed. Peak conditions of each extreme event were noted as the highest combined wave height in a consecutive series above the threshold, and the period and direction of sea or swell, whichever had the highest incremental wave height. Maximum, minimim, mean and standard deviation wave heights also were computed. Each data set included 20 years of record, so the number of extreme events per year (the Poisson lambda parameter) was computed as the total number of events divided by 20.

Initial runs of this extreme event identification program resulted in a

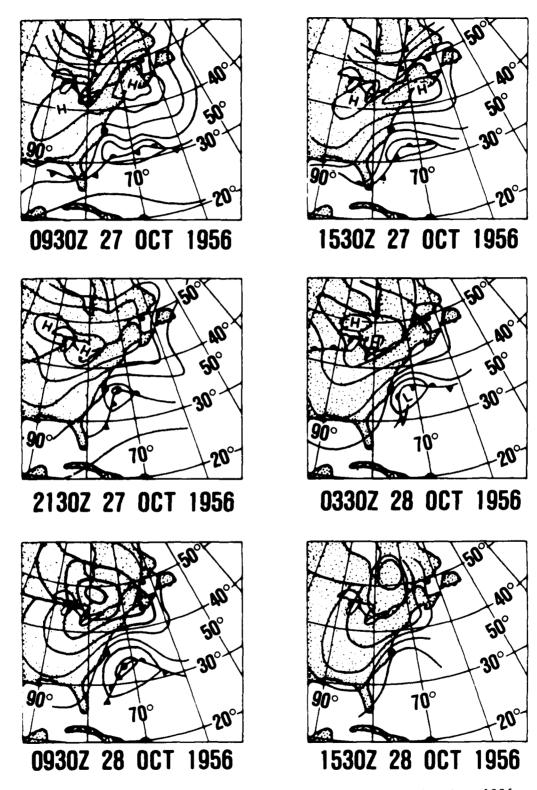


Figure 8. Eastern US surface weather patterns: October 1956

surprisingly large number of extreme events, consistently on the order of 30 to 40 percent of all extreme events identified, to be only 3 hr in duration, i.e., only one record above the threshold. The actual duration could be anywhere from 0 to 6 hr for a single record above the threshold, but an average value of 3 hr was consistently assumed in such cases. Variation of the threshold had little effect on the percentage, although the total number of extreme events was of course affected. Review of climatology considerations inherent in WIS simulations (Kesio and Hayden 1973; and Corson, Resio, and vincent 1980) did not uncover a rationale for excluding a priori durations that short. In fact, a duration of 3 hr is either implicitly or explicitly assumed for peak conditions in many wave forecasts, designs, and research efforts relating to the tradition of sampling wave gages at this interval. Average low pressure systems which would generate extreme wave conditions are known to typically last much longer, however, as in the case of the system illustrated in Figure 8.

In view of this last fact and the example of late October 1956 at Nagshead, the program was adjusted to ignore a lapse below the threshold of only one record (i.e., 6 hr) between consecutive extreme events, as identified previously. This adjustment lowered the number of extreme events of only 3 hr duration (one record above the threshold) only slightly, but a neglect of longer lapses or other adjustments to the identification procedure could not be rationalized. Tables A4-A8 give duration results, following the procedures described above, for the five sites at all thresholds investigated.

Mean and maximum durations for the three Atlantic sites are virtually identical to those reported by Jensen (1983a), with the occasional exception caused by combination of two events separated by only one record with H below the threshold. The mean duration was slightly higher in these few cases.

The percent occurrence of wave heights (percent records H > H1) was of special interest since this statistic for a range of H1 levels is now or will be published and readily available for all WIS stations of all three phases. It was hoped this nondimensional parameter could be used as a tool for choosing threshold levels for duration computations which would preclude many of the iterations which otherwise might be necessary. The number of extreme events per year was also of special interest since this parameter is so important in extremal statistics and expectations.

Figure 9 shows the mean and standard deviation of duration plotted against percent occurrence (actually exceedance) of wave heights above the

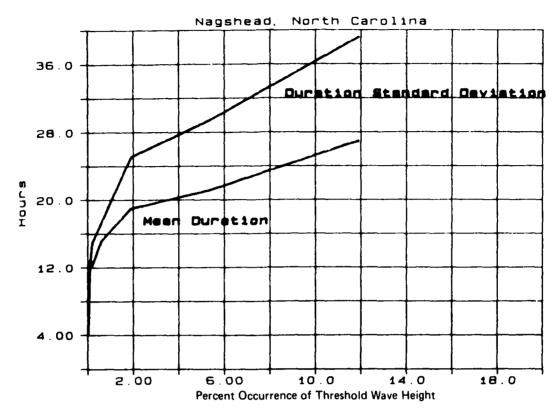


Figure 9. Mean duration and standard deviation versus percent occurrence of wave height threshold, Nagshead, North Carolina

specified threshold for Nagshead. Figures All-Al4 show mean and standard deviation durations for the other four sites plotted against percent occurrence of wave heights above the threshold. Figure 10 shows the nearly linear relationship of the number of extreme events per year with percent occurrence of wave heights above a specified threshold for Nagshead. A similar trend is evident for higher wave with similar percent occurrence at Newport, Oregon, as shown in Figure Al5. These plots in themselves do not indicate an outstanding range of percent occurrence as a choice for definition of extreme events and durations. Some subjective choices can be made since an important purpose of this exercise is to identify extreme events. Clearly, an excessively large number of extreme events per year, say more than 20, will probably include some events that can hardly be regarded as "extremes" in the practical sense. On the other hand, an average number of extreme events per year less than one or two would generally imply exclusion of some events which belong in a

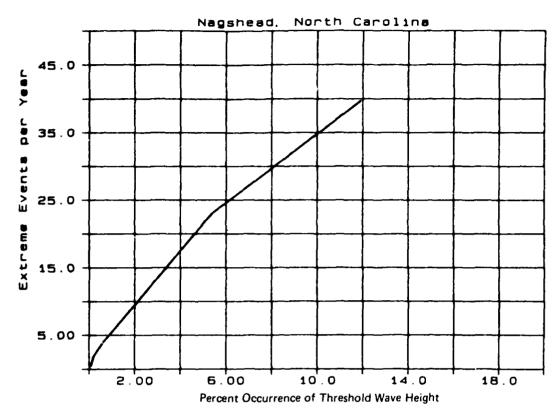


Figure 10. Extreme events per year versus percent occurrence, Nagshead, North Carolina

population of extremes. These considerations are consistent with the author's experience in developing design criteria based on extremal statistics of peak wave height conditions (e.g. Andrew, Smith, and McKee 1985).

A simple linear regression of extreme events per year with percent occurrence of wave heights above the threshold, constrained to pass through the origin, for the 41 cases considered at all five sites indicates that percent occurrence = 0.3λ with a correlation coefficient of 0.97. This relation applies to both the Atlantic and Pacific sites addressed individually, even though the absolute value of wave heights themselves on the Pacific are substantially higher than those on the Atlantic at the same percent occurrence levels. A range in λ of 2 to 20 thus would correspond to a range in percent occurrence of 0.6 to 6.0 percent for the choice of a desirable threshold level, H1. The lower limit of this range would guarantee a sample size of at least 40 extreme events, which is generally desirable for most statistical considerations. The choice of a threshold wave height may be made more precisely when some physical tolerance level is at issue, for example the point

at which some operation at sea must be temporarily terminated.

The other parameters presented in Tables A4-A8 show interesting trends. The minimum duration was 3 hr in every case except one where only three extreme events were identified. A count of extreme events with a 3-hr duration for the Nagshead cases indicated 32 to 48 percent of extreme events shared the minimum duration. No relation of the number of extreme events with a 3-hr duration to the threshold level was apparent. The maximum duration can be seen to be proportional to percent occurrence of wave heights above the threshold and typically many standard deviations above the mean. The mean duration accordingly also is proportional to percent occurrence of wave heights above the threshold. The standard deviation was rarely less than the mean, but always of the same order of magnitude. A lack of central tendency for durations was noted by Houmb and Vik (1977).

Another scheme of extreme event identification was investigated which actually applied a lower threshold Hl in the same way for determination of duration, but only to extreme events whose peak (combined) wave height was above a second higher threshold, H2. The most notable effects of the second threshold were to substantially reduce the number of extreme events per year for a given Hl threshold and to reduce the number of extreme events with a 3-hr duration to zero in nearly every case. Variation of Hl with a fixed H2 had little effect on the number of extreme events per year. The central tendency of durations was somewhat stronger in these subsets, with the standard deviation often, but not always, less than the mean. These two parameters consistently retained the same order of magnitude.

Tables A9 and A10 present the parameter values computed for various combinations of H1 and H2 at the Nagshead and Daytona Beach sites. Figures A16 and A17 show variation of the mean and standard deviation durations with percent occurrence of the lower threshold H1 at a upper threshold H2 tor peak conditions fixed at 300 cm (0.6 percent) and 300 cm (0.8 percent) for the Newport and Nagshead sites. This scheme of double thresholds for extreme event identification was not pursued further since it was considered more desirable to address trends in peak conditions separately from durations above a specified threshold. An approach which addressed marginal distributions versus conditional distributions was preferred.

CHAPTER VII: DISTRIBUTION OF DURATIONS

Method of Analysis

The cumulative probability of durations derived by the single threshold method described above is estimated by application of a plotting formula commonly applied in analyses of this type (Gumbel 1958, and Isaacson and MacKensie 1981):

$$F(t_i) = \frac{i}{(n+1)}$$
 $i = 1, 2, 3, ..., n$ (69)

where $F(t_1)$ is the estimated cumulative probability of the "ith" smallest duration and n is the number of extreme events. Durations are first ordered from smallest to largest for this purpose and the corresponding cumulative probability computed. Other plotting formulae were considered (e.g., Gringorten 1963), but this more commonly used approach is preferable for general application since no additional parameters need be estimated.

Two continuous distributions are considered as models for the cumulative probability of durations because of their common application to peak wave height conditions: the Extremal (Fisher-Tippett) Type I and the Weibull distributions. An existing FORTRAN program (US Army Engineer Waterways Experiment Station 1985), originally designed to fit these distributions to wave height data by the method of least squares, applying the plotting formula convention described above, was adapted to work instead with duration data derived by the extreme event identification program. The extreme event identification program was ultimately combined with the program-estimating distribution parameters and titled STRMDIST, a listing of which is presented in Appendix C. The program STRMDIST, in addition to the extreme event indentification and duration derivation computations already described, computes distribution parameters (ϵ and β for the Extremal Type I and α and β for the Weibull), estimated (distribution) mean and standard deviation, correlation coefficient, sum of the square residuals, and standard error. These parameters also are computed for peak wave heights of extreme events identified. Tables All-Al5 give results of the STRMDIST analysis for five Phase III sites.

Discussion of the Distribution Analysis

Figures 11 and 12 demonstrate fit of the least squares regression distribution to the data as represented by the plotting formula for one case each

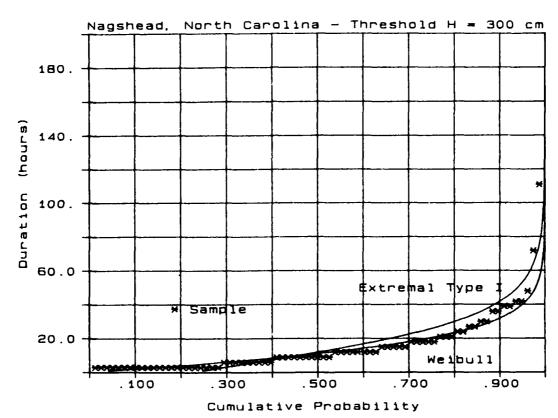


Figure 11. Duration cumulative probability: Nagshead, North Carolina

at Nagshead, North Carolina, and Newport, Oregon. The Weibull distribution in both these cases can be seen to generally fit the overall data spread better, but the Extremal Type I comes closer to the few most extreme durations. The correlation coefficients, sums of square residuals, and standard errors in Tables All-Al5 indicate that the Weibull distribution generally fits the data better than the Extremal Type I, but both distributions fit it acceptably well in practical terms. Correlation coefficients above 0.90 would provide a rule-of-thumb acceptable fit in exercises of this type with weather-related data. Both distributions generally exceed this criterion.

Figure 13 shows correlation coefficients for both distributions plotted against percent occurrence of wave heights above the specified threshold, H1, for Nagshead. Figure A16 shows a similar plot for Newport.

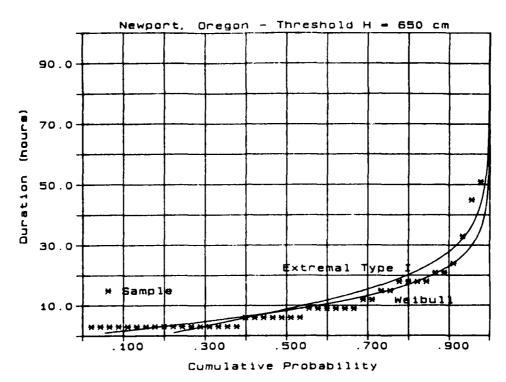


Figure 12. Duration cumulative probability: Newport, Oregon

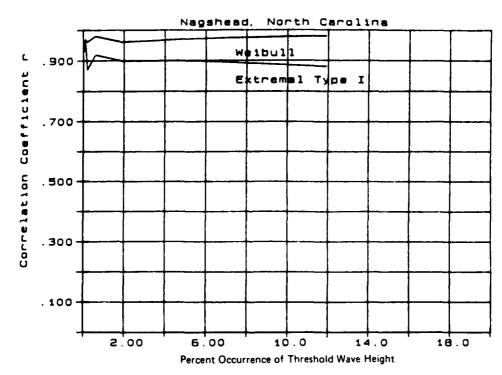


Figure 13. Correlation coefficient versus percent occurrence of wave height threshold, Nagshead, North Carolina

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Figures A17 and A18 show the correlation coefficients for both distributions at Nagshead and Newport plotted against the number of extreme events per year. No obvious maximum occurs which could reliably be taken as an indication of an optimal choice for either λ or H1.

Figures A19 and A21 are plots of the standard error against extreme events per year and percent occurrence for Nagshead, North Carolina. Figures A20 and A22 show the same information for Newport, Oregon. Again, no obvious minimum generally occurs to indicate an optimal choice for λ or H1.

Figures 14 and 15 are graphs of the sample and distribution means and sample and distribution standard deviations plotted against percent occurrence and the number of extreme events per year both for Nagshead. Figures A23 and A24 are similar graphs for Newport. The Extremal Type I distribution mean and standard deviation can be seen to generally come closer to the sample mean and standard deviation. This is desirable, particularly in the case of the mean. The Central Limit Theorem states that sample means from an infinite population can be considered as random variables with a mean equal to the population mean. The standard deviation, as a measure of the spread of duration values about the mean, is an important indicator of how conservative a parameterized distribution might be. The Extremal Type I distribution can be seen to be closer to and consistently larger than (i.e. on the conservative side of) the sample standard deviation. The Weibull distribution standard deviation is both farther from the sample standard deviation and generally lower, i.e., predicting more central tendency than the sample. The Extremal Type I distribution in these respects appears superior to the Weibull distribution.

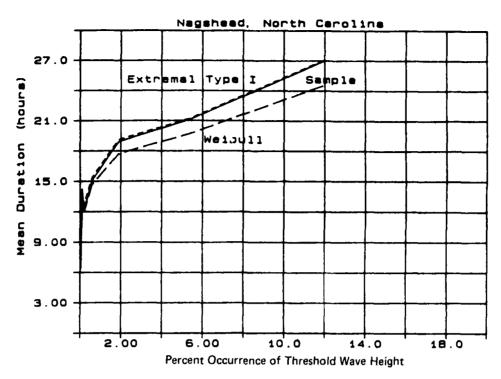


Figure 14. Mean duration versus percent occurrence of wave height threshold, Nagshead, North Carolina

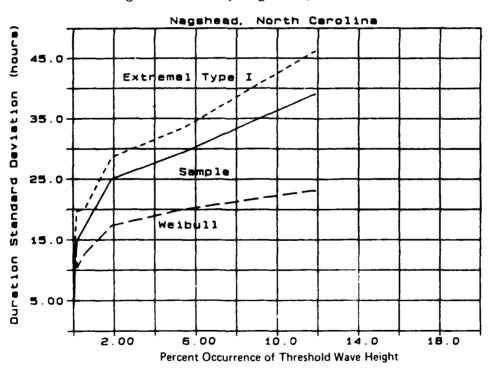


Figure 15. Duration standard deviation versus percent occurrence of wave height threshold, Nagshead,
North Carolina

CHAPTER VIII: RELATIONSHIP OF DURATION TO PEAK CONDITIONS

Method of Analysis

The potential linear or nonlinear relationship of an extreme event's duration with peak conditions of the extreme event were investigated with the aid of statistical software package SPSS (Nie et al. 1975). The stepwise multiple regression capabilities of SPSS were of particular value in testing whether extreme event duration appeared to be dependent on peak conditions, as measured by various parameters such as H , T , H^2 , T^2 , H/gT^2 , H^2L , and direction. Simple linear regressions of extreme event durations, as derived by a range of thresholds, first were performed. In the same program execution, SPSS allowed a stepwise multiple regression of duration against H , H^2 , T , and T^2 to be performed. This procedure estimated the incremental contribution of each of these potentially controlling (independent) variables to the data fit by the least squares method. An equation of the following form was thus possible, assuming the contribution of each of these tested parameters was significant:

$$t = aH + bH^2 + cT + dT^2$$
 (70)

where a, b, c, and d are constants.

The purpose in this exercise was not to derive a predictor equation, but to see if a significant relationship existed. Therefore, the obvious interdependence of H^2 with H and T^2 with T was not of undesirable consequence. One common technique to test for existence of a nonlinear relationship, versus a linear relationship, is also to test the square of the variable on a trial basis. A substantially improved fit with the square of the parameter included in the regression equation generally indicates that a nonlinear relationship, whether polynomial or otherwise, is more reliable than a simple linear relationship. The correlation coefficient, r, as applied above in the fit of distribution functions, was taken as the primary measure of the strength of a relationship in this analysis.

Tables A16-A20 show results from execution of SPSS for all cases tested for each of the five Phase III sites. A listing of the SPSS command file used to perform each of these executions is presented in Appendix D along with a

sample output. The tables give correlation coefficients for duration against H , $\rm H^2$, T , and T (individual simple linear regressions), against all four of these parameters in a stepwise procedure and against $\rm H^2L$ (simple linear regression). There is little indication in any case of a linear relationship of duration with $\rm H/gT^2$ with correlation coefficients for this parameter consistently near zero. Similarly, correlation coefficients of duration with predominant direction of wave propagation at the peaks of extreme events were consistently near zero.

Discussion of Results of the Regression Analysis

The parameter H , the peak zero moment wave height, is consistently the most significant parameter, which confirms that an extreme event identification and duration definition procedure using this parameter is best. Another notable trend indicated by the above results is the observation that correlation coefficients for the Pacific sites are consistently lower than those for Atlantic sites. A possible explanation for this is the fact that Pacific storms typically form well away from the coast and travel onshore. They tend to be well formed when their effects first become significant and their tracks are more or less in the same direction (eastward to some degree). Atlantic (extratropical) storms can form onshore and travel seaward, travel longshore, or linger in one spot, as exemplified by the previous account of conditions in late October 1956. This more variable track (particularly the potential for a roughly stationary storm) may cause the duration above a specified threshold in many cases to be more dependent on the time-history of the storm's internal intensity than its track past a fixed site.

There was no strong correlation of duration (applying the rule-of-thumb criterion of 0.90) with any of the variables on either coast. The regression slopes, i.e. the β parameter in Equation 40, also were consistently small numbers, much closer to zero than to one. The low slopes, even for H , indicate that dependence of durations on peak wave conditions is weak. A fully rigorous proof of dependence or otherwise would require many more tests and computations than those presented here. The lack of an obvious strong dependence, however, raises the suggestion that, for practical purposes, extreme event duration might be taken as independent of peak conditions of the extreme event. This would make estimates of joint probability, for example forecast

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of durations of wave heights above a threshold for a rare event (e.g., the 50-or 100-year extreme event), relatively easy to compute. An example of how such an estimate might be made follows:

Example Computation of Peak Wave Height and Duration Joint Probability

Problem: What is the joint probability of zero moment wave heights greater than 3.0 m lasting longer than 12 hr during an extreme event whose peak zero moment wave height is greater than 4.5 m at Nagshead, North Carolina?

Solution: The definition of duration at Hl allows the associated parameters presented in Table Al2 to be applied. Choosing the Extremal Type I distribution to represent both marginal distribution of peak wave heights and marginal distribution of durations: $\epsilon_{\rm t} = 6.30$, $\beta_{\rm t} = 15.8$, $\epsilon_{\rm H} = 326.3$, and $\beta_{\rm H} = 48.0$. The Poisson parameter, λ , from Table A5, is 3.8. The marginal

probabilities of exceedance are:

$$P(t' > t) = 1 - F(t) = 1 - \exp \left\{-\exp\left[-\frac{(t - \epsilon_t)}{\beta_t}\right]\right\}$$
$$= 1 - \exp\left\{-\exp\left[-\frac{(12 - 6.30)}{15.8}\right]\right\}$$

= 0.502

$$P(H' > H) = 1 - F(H) = 1 - \exp \left\{-\exp \left[-\frac{(H - \epsilon_H)}{\beta_H}\right]\right\}$$

$$= 1 - \exp \left\{-\exp \left[-\frac{(450 - 326.3)}{48.0}\right]\right\}$$

$$= 0.073$$

The joint probability, taken as the product of independent marginal probabilities defined from the same population (HI = 300), is:

$$P(t' > 12, H' > 450 \mid H1 = 300) = 0.502(0.073) = 0.037$$

The associated return period is:

RT(t,H) =
$$\frac{1}{\{\lambda[1-F(t,H)]\}} = \frac{1}{[3.8(0.037)]} = 7.0 \text{ years}$$

The associated nonencounter probability in a 50-year time period is:

NE(t,H) =
$$\exp\left[\frac{-L}{RT(t,H)}\right] = \exp\left(\frac{-50}{7.0}\right) = 0.00079 = 0.08%$$

The associated risk of encountering such a condition in a 50-year time span is 1 - NE(t, H) = 0.921 = 92.1%.

Discussion: Given the assumptions stated above, the probability of exceedance of a peak wave height of 4.5 m of any duration is 7.3 percent. The condition of duration exceeding 12 hr eliminates about half of the possibilities; therefore, the joint probability is about half as much. The joint return period is also correspondingly longer. The Poisson assumption inherent in definition of return period and nonencounter probability can be extended to the joint peak wave height and duration distribution if waiting periods between extreme events are much greater than durations of the extreme events. The Poisson distribution is a discrete distribution, and its application technically extends only to discrete events.

CHAPTER IX: CONCLUSIONS

Literature Review

A review of scientific and engineering literature related to duration of sea states reveals little direct work in this area. The work of Houmb and Vik (1977) is most pertinent to objectives of this study. These investigators worked with several years of intermittently measured wave information at five points along the North Sea coast of Norway. They found the duration of extreme sea states, as defined by the exceedance of a wave height threshold, to fit a Weibull distribution. They approached the problem as much as possible from a theoretical perspective in order to maximize the reliability of observations based on limited data.

Identification of Extreme Events

This study applies the Phase III (shallow water) Wave Information Studies (WIS) database of hindcast wave data because of its unusually long, continuous 20-year period of record and because of its synoptic (ocean wide) perspective on wave conditions. The WIS numerical simulations involve some practical simplifications, but no database of measured wave information is available which could be used to investigate such a long period of record over a wide geographical area. Data from five Phase III stations are applied in this study to investigate duration of extreme wave conditions. Three are on the Atlantic coast (from New Jersey to central Florida) and two are on the Pacific coast (Oregon to central California).

The conventional parameter for long-term wave statistics, zero moment wave height, is chosen as the most practical and reliable indicator of intensity of wave conditions. A computer program is presented which reviews

Phase III information and records the number of sequential records (each 3 hr apart) in which the geometric average (combined) sea and swell wave height is above a specific threshold. A single record below the threshold between two that were above is ignored, i.e., the two records above are treated as part of a single event. The percent occurrence of waves above a threshold is found to vary linearly with the number of extreme events identified, regardless of absolute intensity of wave climate on either coast.

Distribution of Durations

The Weibull and Extremal Type I distributions are fit by the method of least squares to durations of extreme events identified and to peak wave heights. Both distributions show acceptable correlation to the wave data, but the Extremal Type I is found to provide superior estimates of both durations and peak wave heights.

Relationship of Duration to Peak Intensity

A multilinear regression analysis is performed to address the potential relationship of extreme event duration to peak conditions of the extreme event. Peak intensity, as measured by the zero moment wave height, has only a weak linear relationship to duration. Other alternate parameters of intensity show little evidence of significant linear relation to duration. The investigation does not rigorously prove statistical independence, but the assumption of independence of duration from peak intensity is proposed as an expedient measure. This assumption greatly simplifies prediction of durations of wave conditions above a critical threshold.

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ADDITIONAL FIGURES AND TABLES

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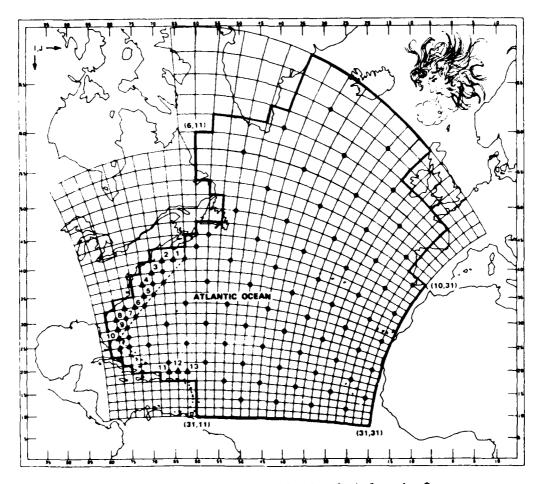


Figure Al. WIS Phase I grid, North Atlantic Ocean

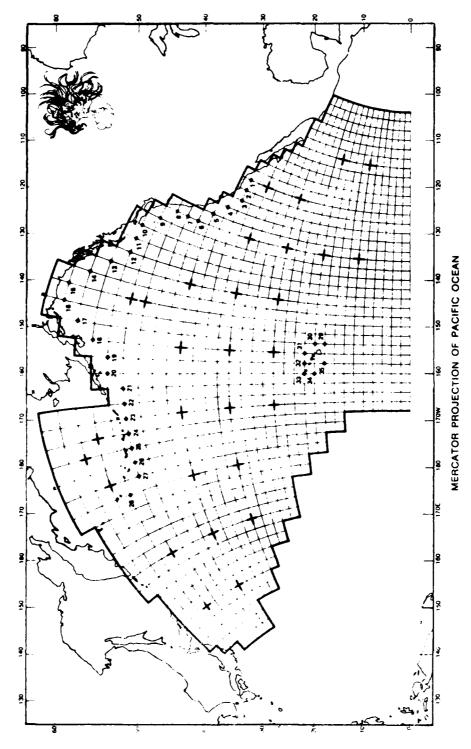


Figure A2. WIS Phase I grid, North Pacific Ocean

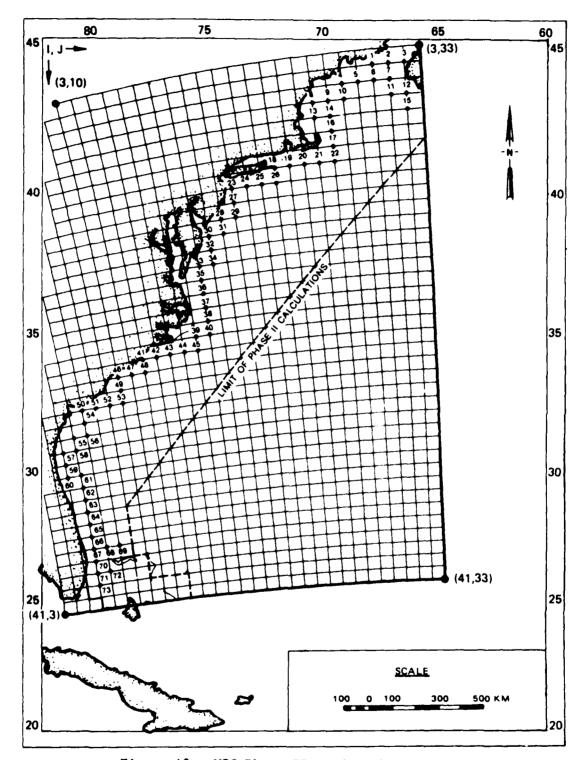


Figure A3. WIS Phase II grid, Atlantic coast

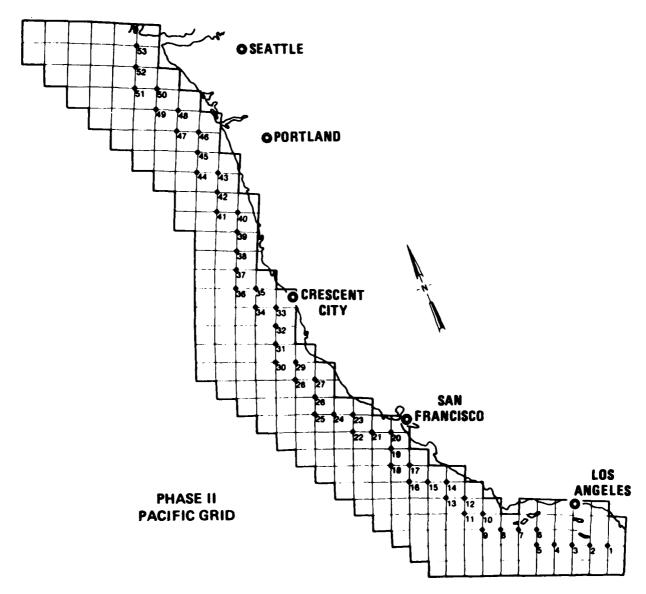


Figure A4. WIS Phase II grid, Pacific coast

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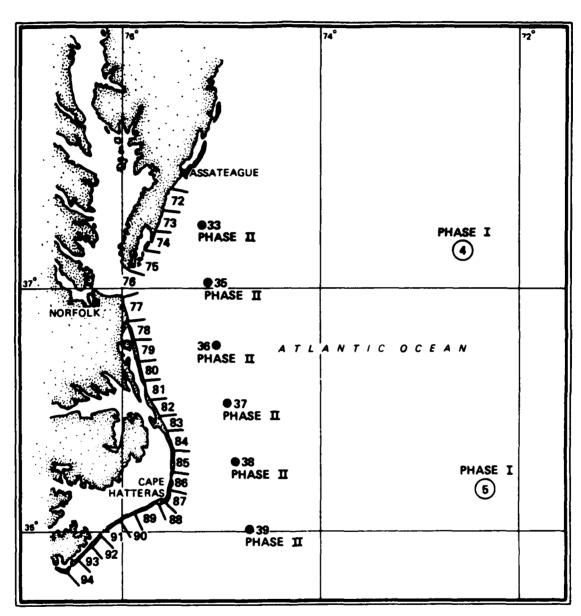


Figure A5. Mid-Atlantic coast portion, WIS Phase III stations

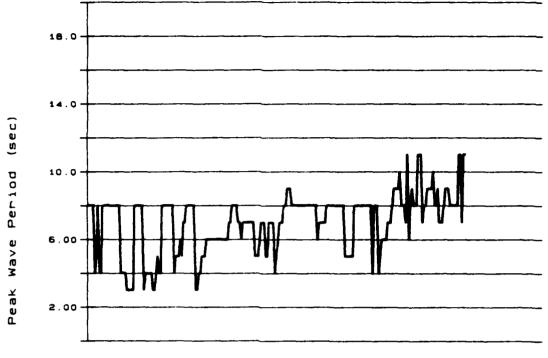


Figure A6. Wave period time series, Nagshead, NC, October 1956

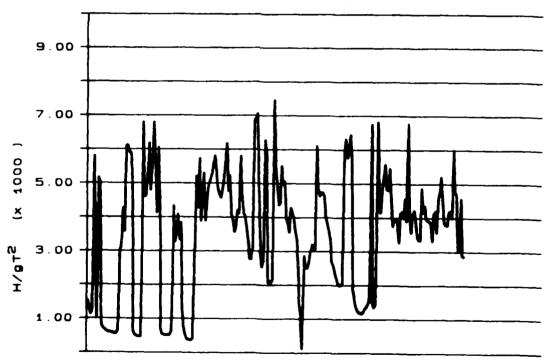


Figure A7. Wave steepness time series, Nagshead, NC, October 1956

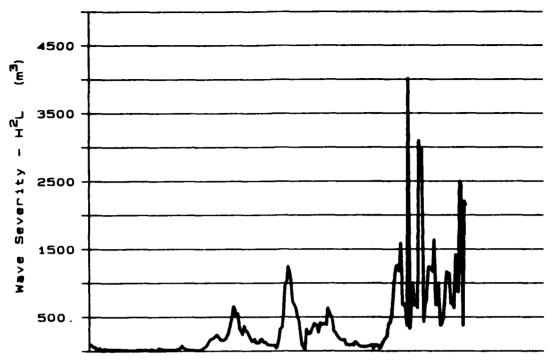


Figure A8. Wave severity time series, Nagshead, NC, October 1956, plotted as $\mathrm{H}^2\mathrm{L}$

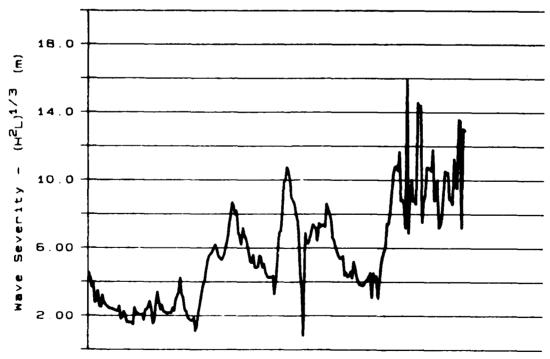


Figure A9. Wave severity time series, Nagshead, NC, October 1956, plotted as $(\mathrm{H}^2\mathrm{L})^{1/3}$

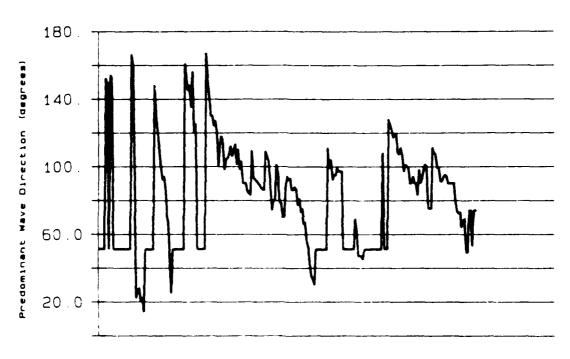


Figure AlO. Wave direction time series, Nagshead, NC, October 1956

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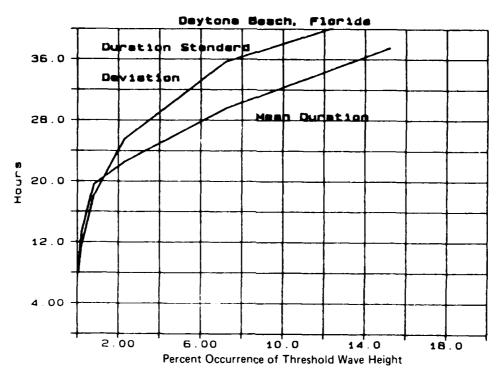


Figure All. t and st versus percent occurrence, $$\operatorname{\textsc{Daytona}}$$ Beach, FL

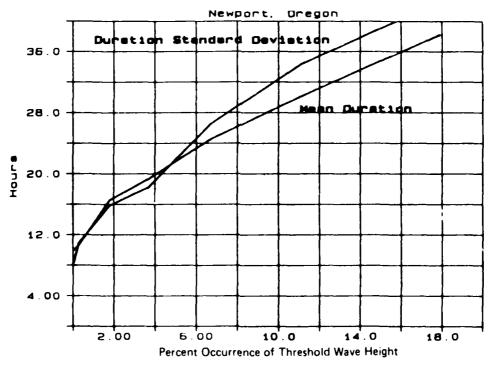


Figure A12. t and st versus percent occurrence, Newport, OR

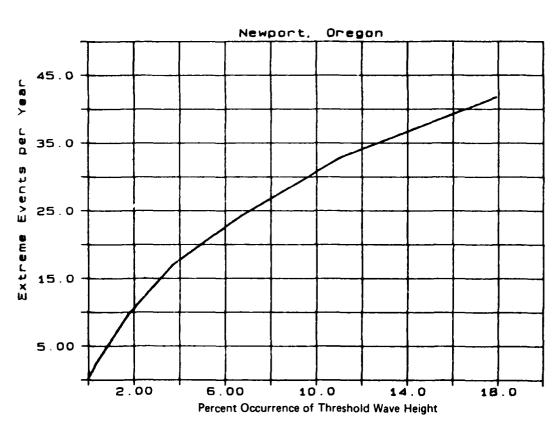


Figure Al3. Extreme events per year versus percent occurrence, Newport, OR

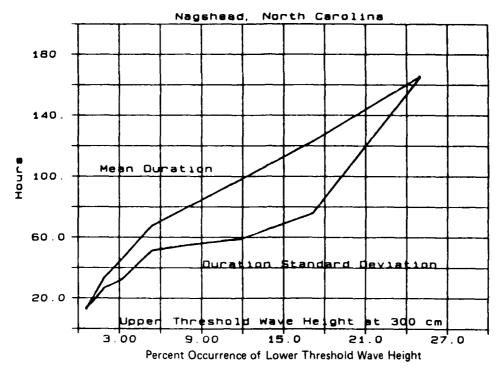


Figure Al4. t and st versus percent occurrence with peak threshold, Nagshead, NC

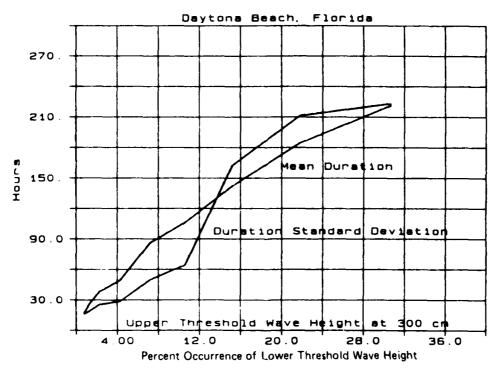


Figure Al5. t and st versus percent occurrence with peak threshold, Daytona Beach, ${\rm FL}$

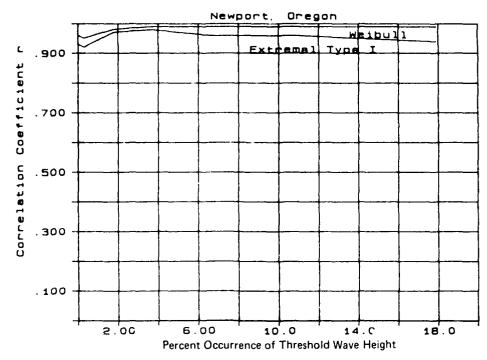


Figure Al6. Correlation coefficient versus percent occurrence, Newport, OR

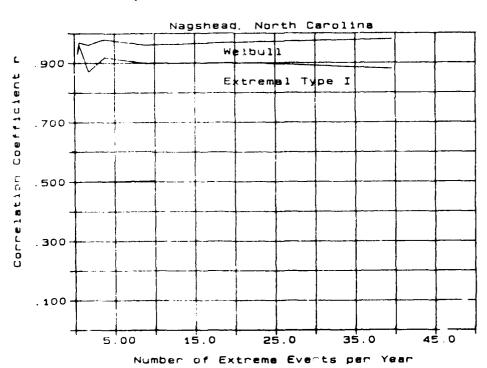


Figure Al7. Correlation coefficient versus extreme events per year, Nagshead, NC

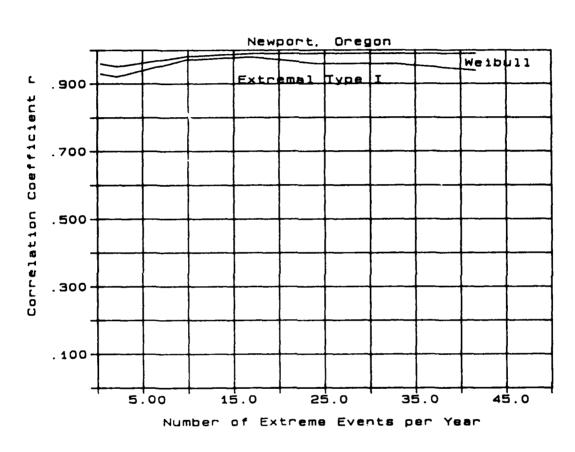


Figure Al8. Correlation coefficient versus extreme events per year, Newport, OR

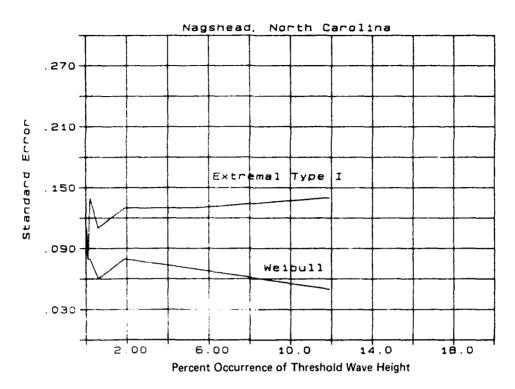


Figure A19. Standard error versus percent occurrence, Nagshead, NC

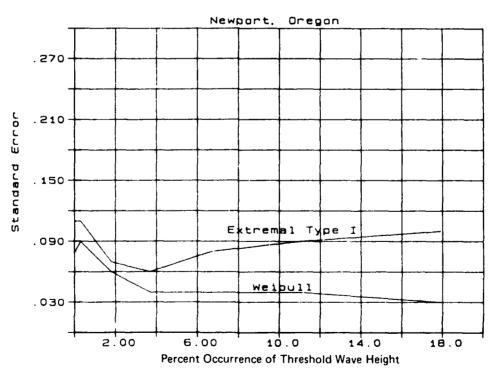


Figure A20. Standard error versus percent occurrence, Newport, OR

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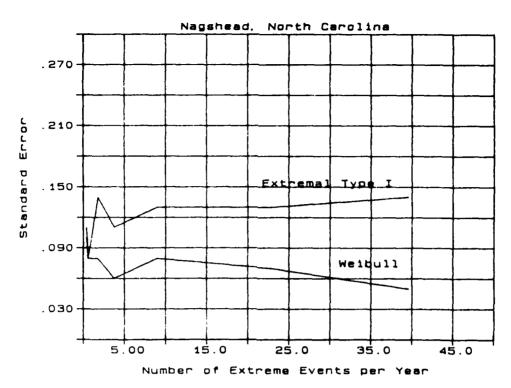


Figure A21. Standard error versus extreme events per year, Nagshead, NC

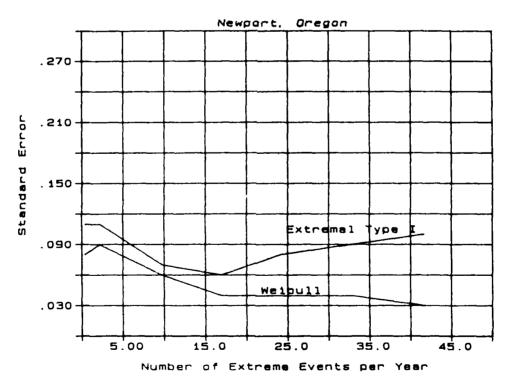


Figure A22. Standard error versus extreme events per year, Newport, OR

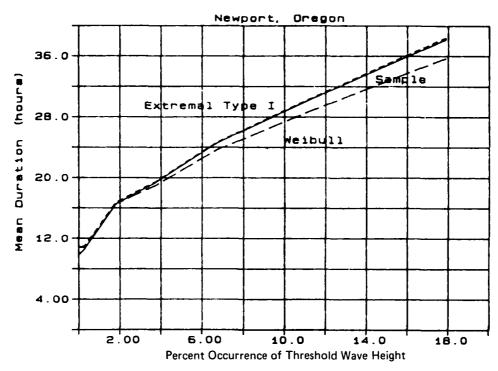


Figure A23. Mean duration versus percent occurrence, Newport, OR

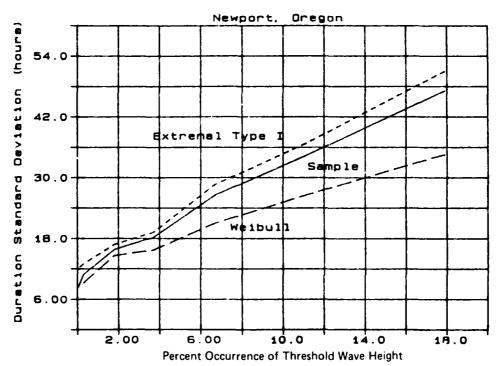


Figure A24. Duration standard deviation versus percent occurrence, Newport, OR

Table A1 October 1956 Phase III Data, Nagshead, NC

Station: A3083	A3083	Sea	a Readings-	{S}	WS	all Readir	1gs		-Combined-	
Date		Height	Period	Direct	Height	Period	Direct		Period	Direct
YY/MM/DD	Hour	(cm)	(secs)	(azim)	(cm)	(cm) (secs) (azim)	(azim)	(E)	(secs)	(azim)
56/10/27	00:00	219	7	06	0	0	0	219	7	96
56/10/27	03:00	137	9	103	144	11	83	462	:	83
56/10/27	00:90	208	9	86	119	11	83	240	9	98
56/10/27	00:60	261	8	06	0	0	0	261	æ	96
56/10/27	12:00	282	6	93	0	0	0	282	6	93
56/10/27	15:00	258	80	98	0	0	0	258	∞	98
56/10/27	18:00	592	80	101	0	0	0	592	ထ	101
56/10/27	21:00	250	80	100	0	0	0	250	ω	100
56/10/28	00:00	93	5	130	395	11	75	901	11	75
56/10/28	03:00	117	5	130	378		75	396	11	75
56/10/28	00:90	149	9	127	369	=	75	398	1	75
56/10/28	00:60	234	7	111	0	0	0	234	7	111
56/10/28	12:00	260	ω	108	0	0	0	260	∞	108
56/10/28	15:00	273	80	107	0	0	0	273	∞	107
56/10/28	18:00	314	6	66	0	0	0	314	6	66
56/10/28 21:00	21:00	311	6	66	0	0	0	311	6	66

Table A2 October 1956 Phase II Data, Nagshead, NC

Station:	A2037	Sea	ea Readings-	3	MS	Swell Readings	1gs		Combined	
Date YY/MM/DD	Hour	Height (cm)	Period (secs)	Direct (azim)	Height (cm)	Period (secs)	Direct (azim)	Height (cm)	Period (secs)	Direct (azim)
56/10/27	00:00	488	7	63	0	0	0	488	7	63
56/10/27	03:00	167	9	611	438	=	77	6917	Ξ	77
56/10/27	00:90	111	9	45	119	11	78	492	9	45
56/10/27	00:60	508	ω	62	0	0	0	508	œ	62
56/10/27	12:00	204	6	69	0	0	0	507	6	59
56/10/27	15:00	98ħ	œ	715	0	0	0	486	œ	717
56/10/27	18:00	473	∞	817	0	0	0	473	œ	817
56/10/27	21:00	₄ 52	∞	617	0	0	0	452	ω	617
56/10/28	00:00	110	2	17	399	11	82	717	11	82
56/10/28	03:00	139	2	17	382	11	85	407	-	82
56/10/28	00:90	197	9	17	373	11	83	422	11	83
56/10/28	00:60	944	7	36	0	0	0	944	7	36
56/10/28	12:00	425	80	37	0	0	0	425	œ	37
56/10/28	15:00	447	ω	39	0	0	0	L th th	∞	39
56/10/28	18:00	117	6	617	0	0	0	117	6	617
56/10/28	21:00	505	6	20	0	0	0	505	6	20

Table A3 October 1956 Phase I Deepwater Data, Offshore of Cape Natteras, NC

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Table A4

Duration Information for Atlantic City, NJ

H1	H > H1	H > H1	Number of Events	Number of Events/yr	t _{min}	t _{max}	t hrs	σ _t
200	1442	2.5	323	16.2	3	54	12.8	10.9
250	384	0.7	112	5.6	3	30	9.9	7.3
300	81	0.1	29	1.4	3	24	8.1	6.8
350	18	0.03	9	0.4	3	15	5.7	4.1

Table A5

Duration Information for Nagshead, NC

H1 cm	H > H1	H > H1	Number of Events	Number of Events/yr	t _{min}	t _{max}	t hrs	σ _t
150	6983	11.9	792	39.6	3	570	26.9	39.1
200	3167	5.4	460	23.0	3	306	21.1	29.4
250	1093	1.9	179	9.0	3	165	18.9	25.0
300	374	0.6	77	3.8	3	111	15.1	17.2
350	143	0.2	36	1.8	3	84	12.2	14.9
400	56	0.10	13	0.6	3	42	12.9	12.0
450	16	0.03	8	0.4	3	15	6.4	4.1

Table A6

Duration Information for Daytona Beach, FL

H1 cm	H > H1	H > H1	Number of Events	Number of Events/yr	t _{min}	t _{max}	t hrs	o _t
150	8855	15.2	716	35.8	3	1032	37.5	65.5
200	4183	7.2	432	21.6	3	303	29.5	35.6
250	1340	2.3	186	9.3	3	129	22.5	25.5
300	478	0.8	75	3.8	3	81	19.5	18.0
350	143	0.2	33	1.6	3	60	13.3	11.5
400	31	0.05	12	0.6	3	33	8.0	8.4
450	8	0.01	3	0.2	3	18	9.0	7.9

Table A7

Duration Information for Newport, OR

H1 cm	H > H1	H > H1	Number of Events	Number of Events/yr	t _{min}	t _{max}	t hrs	σ _t
400	10472	17.9	834	41.7	3	405	38.2	47.1
450	6494	11.1	658	32.9	3	279	30.1	34.3
500	3897	6.7	484	24.2	3	261	24.5	26.5
550	2152	3.7	341	17.0	3	108	19.3	18.2
600	1049	1.8	196	9.8	3	81	16.5	15.8
650	151	0.3	44	2.2	3	51	10.6	10.9
700	22	0.04	7	0.4	3	27	9.9	8.3

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Table A8

Duration Information for Half-Moon Bay, CA

H1 em	H > H1	H > H1	Number of Events	Number of Events/yr	t _{min} hrs	t _{max}	t hrs	σ _t
500	768	1.3	105	5.2	3	123	21.7	21.2
550	373	0.6	50	2.5	3	108	22.0	21.7
600	168	0.3	23	1.2	3	78	21.5	17.8
650	17	0.03	3	0.2	12	21	17.0	4.6

Duration Parameters with a Peak Wave Height Threshold,
Nagshead, NC

H1 cm	H2 cm	% Records H > H1	Number of Events	Number of Events/yr	t _{min}	t _{max}	ŧ hrs	^o t hrs
100	300	25.0	52	2.6	18	1056	165.3	164.8
125	300	17.1	52	2.6	15	381	122.6	75.5
150	300	11.9	52	2.6	12	375	98.0	58.8
175	300	8.4	54	2.7	9	333	81.7	55.3
200	300	5.4	55	2.8	6	306	67.6	51.3
225	300	3.2	62	3.1	6	174	45.6	32.1
250	300	1.9	68	3.4	3	165	33.1	26.5
275	300	1.1	77	3.8	3	117	20.5	17.6
300	300	0.6	90	4.5	3	102	12.5	13.6
100	250	25.0	118	5.9	18	1056	119.6	120.6
125	250	17.1	119	6.0	15	381	92.5	64.1
150	250	11.9	123	6.2	12	375	74.6	52.6
175	250	8.4	129	6.4	6	333	60.0	47.4
200	250	5.4	136	6.8	3	306	45.4	40.3
225	∠5 0	3.2	173	8.6	3	174	26.0	25.6
250	250	1.9	212	10.6	3	165	15.5	19.9

Duration Parameters with a Peak Wave Height
Threshold, Daytona Beach, FL

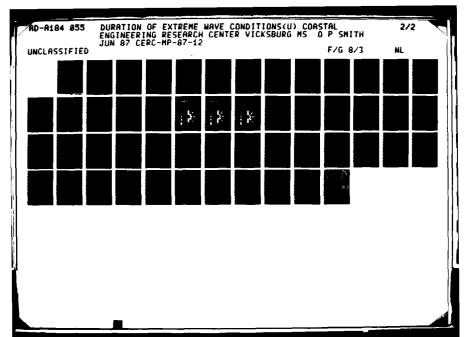
H1 cm	H2	% Records H > H1	Number of Events	Number of Events/yr	t _{min}	t _{max}	t hrs	σ _t
100	300	30.7	49	2.4	36	1197	221.0	223.1
125	300	21.8	49	2.4	24	1191	184.3	211.1
150	300	15.2	51	2.6	15	1035	141.8	162.1
175	300	10.6	54	2.7	12	354	105.7	63.8
200	300	7.2	54	2.7	9	303	85.8	49.5
225	300	4.3	68	3.4	3	141	49.1	28.4
250	300	2.3	69	3.4	3	114	38.1	25.1
275	300	1.3	75	3.8	3	87	25.5	18.7
300	300	0.8	84	4.2	3	81	17.1	15.8
100	250	30.7	119	6.0	9	1197	162.6	162.1
125	250	21.8	121	6.0	9	1191	131.0	147.8
150	250	15.2	127	6.4	6	1035	101.5	111.5
175	250	10.6	133	6.6	6	354	78.0	52.4
200	250	7.2	140	7.0	3	303	59.9	41.3
225	250	4.3	193	9.6	3	141	30.9	24.8
250	250	2.3	238	11.9	3	114	16.9	20.0

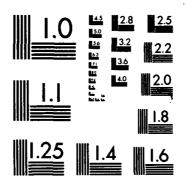
Table A11

Distribution Parameters for Durations and Peak Wave Heights at Atlantic City, NJ

		-Duratio		Pea	k Wave H	eight
Parameter	Sample	Type I	Weibull	Sample	Type I	Weibull
H1 = 200 cm (2.5%)	occurre	nce leve	1)			
ε/a	-	7.75	1.23	-	222.6	7.06
8	-	8.93	14.0	-	32.1	258.0
x	12.8	12.9	13.1	240.9	241.1	241.4
σ	10.9	11.4	10.7	39.5	41.1	40.2
r	-	0.97	0.97	-	0.98	0.93
Eres ^{2*}	-	1.40	1.47	-	0.838	3.51
std.err.	-	0.066	0.068	-	0.051	0.104
H1 = 250 cm (0.7%)	occurre	nce)				
ε/a	-	6.36	1.45	-	271.4	9.77
В	-	6.23	11.04	-	27.3	301.9
x	9.9	10.0	10.0	286.8	287.2	287.0
σ	7.3	8.0	7.0	32.7	35.0	35.3
r	-	0.97	0.97	-	0.99	0.93
Eres ^{2#}	•	0.512	0.534	-	0.222	1.21
std.err.	-	0.068	0.070	-	0.045	0.105

^{*}Sum of the square residuals.





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Table A11 (Concluded)

	Duration				Peak Wave Height		
<u>Parameter</u>	<u>Sample</u>	Type I	Weibull	<u>Sample</u>	Type I Weibull		
H1 = 300 cm (0.1% occurrence)							
ε/α	-	4.65	1.45	-	318.7 12.6		
В	-	6.39	9.20	-	25.2 345.5		
x	8.1	8.3	8.6	332.2	333.2 331.6		
σ	6.8	8.2	7.1	27.7	32.3 32.0		
r	•	0.92	0.92	-	0.98 0.97		
Ires ^{2#}	-	0.366	0.366	-	0.067 0.139		
std.err.	-	0.116	0.116	-	0.050 0.072		
H1 = 350 cm (0.03	s occurr	ence)					
ε/α	-	3.43	1.36	-	354.5 15.4		
В	-	4.57	6.78	-	21.8 376.6		
x	5.7	6.1	6.2	365.2	367.1 364.0		
σ	4.1	5.9	4.6	18.4	28.0 29.0		
r	-	0.89	0.89	-	0.86 0.77		
Eres ^{2*}	-	0.128	0.122	-	0.159 0.243		
std.err.	-	0.135	0.132	-	0 .151 0.186		

[•] Sum of the square residuals.

Table A12

<u>Distribution Parameters for Durations and Peak Wave Heights at Nagshead, NC</u>

	Duration			Pea	Peak Wave Height			
Parameter	Sample	Type I	<u>Weibull</u>	<u>Sample</u>	Type I	Weibull		
H1 = 150 cm (11.9% occurrence)								
ε/α	•	6.19	1.06	-	181.5	4.80		
ß	-	36.0	25.0	-	44.5	226.8		
x	26.9	27.0	24.5	207.0	207.2	207.7		
σ	39.1	46.2	23.2	55.1	57.1	49.4		
r	-	0.88	0.98	-	0.99	0.95		
Σres ²	-	14.5	2.18	-	1.59	6.43		
std.err.	-	0.14	0.05	-	0.04	0.09		
H1 = 200 cm (5.4	% occurre	ence)						
ε/α	-	6.20	0.99	-	218.5	5.55		
ß	-	26.0	19.6	-	43.2	264.8		
×	21.1	21.2	19.7	243.2	243.4	244.6		
σ	29.4	33.4	20.0	51.8	55.4	51.0		
r	-	0.90	0.97	-	0.95	0.89		
Σres ²	-	7.35	2.53	-	3.50	7.81		
std.err.	-	0.13	0.07	-	0.09	0.13		

(Sheet 1 of 4)

Table A12 (Continued)

	Duration				Peak Wave Height			
<u>Parameter</u>	<u>Sample</u>	Type I	Weibull		mple		Weibull	
H1 = 250 cm (1.9% occurrence)								
ε/α	-	6.21	1.02		-	272.3	6.35	
β	-	22.3	17.9		-	45.4	321.4	
x	18.9	19.1	17.7	29	8.0	298.5	299.1	
σ	25.0	28.6	17.3	5	3.9	58.2	55.0	
r	-	0.90	0.96		-	0.96	0.90	
res ²	-	2.79	1.04		-	1.05	2.73	
std.err.	-	0.13	0.08		-	0.08	0.12	
H1 = 300 cm (0.6%)	occurr	ence)						
ε/α	-	6.30	1.15		-	326.3	7.03	
в	-	15.8	15.5		-	48.0	378.0	
x	15.1	15.4	14.7	35:	3.0	354.0	353.7	
σ	17.2	20.2	12.8	5	5.7	61.6	59.2	
r	-	0.92	0.98		-	0.97	0.92	
Eres ²	-	0.91	0.30		-	0.35	0.95	
std.err.	-	0.11	0.06		•	0.07	0.11	

(Sheet 2 of 4)

Table A12 (Continued)

		-Duratio	n	Peak Wave Height				
<u>Parameter</u>	<u>Sample</u>	Type I	Weibull	Sample Type I Weibull				
H1 = 350 cm (0.2% occurrence)								
ε/α	-	3.9	1.16	- 375.0 7.9				
ß	-	15.2	12.6	- 48.5 426.4				
x	12.2	12.7	12.0	401.2 402.9 401.3				
σ	14.9	19.6	10.4	53.4 62.2 60.3				
r	-	0.87	0.96	- 0.96 0.90				
Σres ²	-	0.68	0.22	- 0.21 0.52				
std.err.	- ,	0.14	0.08	- 0.08 0.12				
H1 = 400 cm (0.10	% occurr	ence)						
ε/a	-	6.8	0.97	- 428.9 8.4				
β	-	12.1	14.3	- 53.5 481.7				
x	12.9	13.8	14.5	456.0 459.8 454.5				
σ	12.0	15.5	15.0	53.4 68.6 64.8				
r	-	0.96	0.97	- 0.97 0.95				
Σres ²	-	0.08	0.06	- 0.05 0.10				
std.err.	-	0.08	0.08	- 0.07 0.10				

(Sheet 3 of 4)

Table A12 (Concluded)

		-Duratio	n	Pea	k Wave He	ight
<u>Parameter</u>	Sample	Type I	Weibull	<u>Sample</u>		Weibull
H1 = 450 cm (0	.03% occurr	ence)				
€/a	-	4.2	1.51	-	460.4	8.6
В	-	4.5	7.52	-	52.9	512.2
x	6.4	6.8	6.8	486.0	490.9	484.1
σ	4.1	5.8	4.6	46.3	67.9	66.9
r	-	0.93	0.94	-	0.89	0.84
Eres ²	-	0.07	0.06	-	0.10	0.16
std.err.	-	0.11	0.10	-	0.13	0.16

Table A13

Distribution Parameters for Durations and Peak Wave Heights at

Daytona Beach, FL

		-Duratio	n	Pea	k Wave Hei	
<u>Parameter</u>	Sample	Type I	Weibull	<u>Sample</u>	Type I	<u>leibull</u>
H1 = 200 cm (7.3)	2% occurre	nce)				
ε/ α	-	12.4	0.96	-	221.3	5.63
в	-	30.0	27.9	-	42.4	266.9
x	29.5	29.7	28.4	295.3	245.7	246.7
σ	35.6	38.5	29.7	51.4	54.4	50.7
r	-	0.95	0.98	-	0.97	0.90
Eres ²	-	3.56	1.11	-	2.43	6.52
std.err.	-	0.09	0.05	-	0.08	0.12
H1 = 250 cm (2.)	3% occurre	ence)				
ε/α	-	10.2	1.01	-	217.0	6.48
В	-	21.6	21.9	-	42.8	318.1
x	22.5	22.7	21.8	295.3	2 95.7	296.3
σ	25.5	27.6	21.7	51.4	54.9	53.5
r	-	0.95	0.98	-	0.97	0.90
Eres ²	-	1.51	0.68	-	1.02	2.79
std.err.	•	0.09	0.06	-	0.07	0.12

Table A13 (Concluded)

				Peak Wave Height
<u>Parameter</u>	<u>Sample</u>	Type I	Weibull	Sample Type I Weibull
H1 = 300 cm (0.8%)	occurre	nce)		
ε/α	-	10.8	1.21	- 329.7 8.43
β	-	15.5	20.6	- 39.4 372.6
x	19.5	19.8	19.3	351.6 352.4 351.7
σ	18.0	19.9	16.1	46.4 50.5 49.7
r	-	0.98	0.99	- 0.99 0.93
Eres ²	-	0.29	0.11	- 0.17 0.78
std.err.	-	0.06	0.04	- 0.05 0.10
H1 = 350 cm (0.2%)	occurre	nce)		
ε/α	•	7.39	1.40	- 375.7 10.4
ß	-	10.9	14.6	- 35.4 413.9
x	13.3	13.7	13.3	394.8 396.2 394.4
σ	11.5	14.0	9.61	39.4 45.4 45.7
r	•	0.95	0.98	- 0.98 0.93
Ires ²	•	0.27	0.10	- 0.13 0.36
std.err.	•	0.09	0.06	- 0.06 0.11

Table A14

Distribution Parameters for Durations and Peak Wave Heights at Newport, OR

		-Duratio		Peak Way				
<u>Parameter</u>	<u>Sample</u>	Type I	Weibull	Sample Type	<u> I Weibull</u>			
H1 = 400 cm (17.9% occurrence)								
ε/α	-	15.4	1.03	- 45	7.45			
ß	-	39.7	36.2	- 6	52.7 527.3			
x	38.2	38.4	35.7	493.9 49	94.1 494.8			
σ	47.1	51.0	34.6	77.7	30.4 78.5			
r	-	0.94	0.99	-	0.99 0.95			
Ires ²	-	8.07	0.87	-	1.52 6.81			
std.err.	-	0.10	0.03	-	0.04 0.09			
H1 = 450 cm (11.1)	% occurr	ence)						
ε/α	-	13.6	1.08	- 50	5.5 7.51			
В	-	28.6	29.4	- 5	53.5 565.4			
x	30.1	30.2	28.6	536.2 53	36.4 536.7			
σ	34.3	36.7	26.5	65.9	63.6 67.7			
r	-	0.96	0.99	-	0.99 0.96			
Ires ²	-	4.75	0.92	-	1.55 3.96			
std.err.	-	0.09	0.04	-	0.05 0.08			

(Sheet 1 of 4)

Table A14 (Continued)

		-Duratio	n	Peak Wave Height
<u>Parameter</u>	Sample	Type I	<u>Weibull</u>	Sample Type I Weibull
H1 = 500 cm (3.7%	occurre	nce)		
e/a	-	11.8	1.13	- 547.4 13.0
8	-	22.3	24.7	- 41.6 594.6
x	24.5	24.6	23.6	571.3 571.5 571.5
σ	26.5	28.6	21.0	51.4 53.4 53.5
r	-	0.96	0.99	- 0.98 0.98
Σres ²	-	3.01	0.72	- 1.24 1.68
std.err.	-	0.08	0.04	- 0.05 0.06
H1 = 550 cm (3.7%	occurre	nce)		
e/a	-	10.8	1.21	- 585.0 18.9
в	-	14.9	20.2	- 29.8 619.3
x	19.3	19.4	18.9	602.0 602.2 602.0
σ	18.2	19.2	15.7	37.1 38.2 39.5
r	-	0.98	0.99	- 0.99 0.97
Ires ²	•	1.32	0.61	- 0.40 1.47
std.err.	-	0.06	0.04	- 0.03 0.07

(Continued)

(Sheet 2 of 4)

Table A14 (Continued)

			n					
<u>Parameter</u>	Sample	Type I	Weibull	Sample Type I Weibull				
H1 = 600 cm (1.8%	occurre	nce)						
ε/α	-	9.1	1.13	- 616.6 24.4				
ß	-	13.1	17.2	- 22.6 643.6				
x	16.5	16.7	16.5	629.5 629.7 629.5				
σ	15.8	16.8	14.6	27.6 29.0 32.1				
r	-	0.97	0.98	- 0.99 0.98				
Σres ²	-	1.09	0.75	- 0.41 2.53				
std.err.	-	0.07	0.06	- 0.05 0.11				
H1 = 650 cm (0.3%	occurre	nce)						
ε/α	-	5.1	1.12	- 663.7 32.2				
ß	-	10.1	11.4	- 18.8 685.1				
x	10.6	11.0	10.8	673.9 674.5 673.5				
σ	10.9	13.0	9.2	21.4 24.1 26.3				
r	-	0.92	0.95	- 0.98 0.93				
Σres ²	-	0.52	0.34	- 0.13 0.50				
std.err.	-	0.11	0.09	- 0.06 0.11				

(Continued)

(Sheet 3 of 4)

Table A14 (Concluded)

 ,		-Duratio	n	Pea	Peak Wave Height				
<u>Parameter</u>	<u>Sample</u>	Type I	Weibull	Sample	Type I	Weibull			
H1 = 700 (0.04\$	occurrenc	e)							
ε/α	-	5.4	1.15	-	711.5	71.9			
В	-	9.4	11.5	-	10.4	721.1			
x	9.9	10.8	10.9	716.4	717.5	715.5			
σ	8.3	12.1	9.5	9.2	13.3	12.6			
r	-	0.93	0.96	-	0.97	0.98			
Eres ²	-	0.06	0.03	-	0.03	0.02			
std.err.	_	0.11	0.08	-	0.07	0.06			

Table A15

Distribution Parameters for Durations and Peak Wave Heights at Half-Moon Bay, CA

		-Duratio	n	Peak Wave Height
Parameter	Sample		Weibull	Sample Type I Weibull
H1 = 500 cm (1.3%	occurre	nce)		
ε/α	-	11.7	1.11	- 533.1 14.2
ß	-	18.0	22.5	- 36.3 574.0
x	21.7	22.0	21.7	553.4 554.0 553.3
σ	21.2	23.1	19.6	43.7 46.6 47.6
r	-	0.98	0.99	- 0.99 0.97
Eres ²	-	0.42	0.19	- 0.12 0.58
std.err.	-	0.06	0.04	- 0.03 0.08
H1 = 550 cm (0.6%	occurre	nce)		
ε/α	-	11.4	1.14	- 575.9 19.4
β	-	19.4	22.9	- 28.3 607.4
×	22.0	22.6	21.8	591.3 592.1 590.8
σ	21.7	24.9	19.2	32.4 36.1 37.7
r	-	0.96	0.99	- 0.99 0.95
res ²	-	0.32	0.09	- 0.06 0.39
std.err.	-	0.08	0.04	- 0.04 0.09

(Continued)

Table A15 (Concluded)

		-Duratio	n	Pea	Peak Wave Height				
Parameter	Sample	Type I	Weibull	Sample	Type I	Weibull			
H1 = 600 cm (0)	.3% occurre	nce)							
ε/α	-	12.8	1.17	-	607.1	20.0			
ß	-	16.6	23.7	-	26.9	637.4			
x	21.5	22.3	22.4	621.3	622.7	620.5			
σ	17.8	21.3	19.3	28.0	34.5	38.4			
r	-	0.99	0.99	•	0.93	0.86			
Eres ²	-	0.05	0.02	-	0.22	0.46			
std.err.	-	0.05	0.03	_	0.10	0.15			

Table A16

Results of Regression of Duration Against Conditions at the Peak of the Event for Atlantic City, NJ

<u>H1</u>	\$(H > H1)	r _H	rH2	r _T	r _T ²	² 2 2 r _{H,H,T,T}	r _{H L}
200	2.5	0.71	0.69	0.54	0.52	0.74	0.66
250	0.7	0.69	0.69	0.40	0.38	0.71	0.63
300	0.1	0.80	0.80	0.23	0.21	0.80	0.65
350	0.03	0.10	0.10	0.24	0.22	0.63	0.21

Results of Regression of Duration Against Conditions at the Peak of the Event for Nagshead, NC

<u>H1</u>	%(H > H1)	r _H	rH_	r _T	r _T ²	2 2 r _{H,H} ,T,T	r _{H L}
200	5.6	0.82	0.81	0.61	0.62	0.82	0.75
250	2.0	0.80	0.79	0.50	0.51	0.81	0.72
300	0.7	0.72	0.72	0.60	0.62	0.74	0.73
350	0.3	0.55	0.54	0.54	0.56	0.70	0.56

Table A18

Results of Regression of Duration Against Conditions at the Peak of the Event for Daytona Beach, FL

<u>H1</u>	\$(H > H1)	r _H	r _H	_r _T	r _T ²	2 2 ^r H,H ,T,T	r _{H L}
200	7.1	0.76	0.73	0.50	0.48	0.79	0.65
250	2.2	0.81	0.79	0.46	0.46	0.82	0.72
300	0.8	0.51	0.50	0.46	0.47	0.59	0.52
350	0.2	0.44	0.46	0.52	0.55	0.66	0.59

Results of Regression of Duration Against Conditions at the Peak of the Event for Newport, OR

<u>H1</u>	\$(H > H1)	r _H	r _H	r _T	r _T 2	2 2 ^r H,H ,T,T	r _{H L}
500	6.9	0.59	0.60	0.32	0.32	0.61	0.55
550	3.8	0.59	0.59	0.24	0.24	0.59	0.50
600	1.9	0.42	0.42	0.20	0.20	0.45	0.38
650	0.3	0.66	0.66	-0.08	-0.08	0.66	0.26

Results of Regression of Duration Against Conditions at the Peak of the Event for Half-Moon Bay, CA

<u>H1</u>	%(H > H1)	r _H	r _H	r _T	r _T ²	2 2 ^r H,H ,T,T	r _{H L}
500	1.3	0.62	υ .6 1	0.24	0.24	0.63	0.52
550	0.6	0.52	0.52	0.22	0.22	0.57	0.48
600	0.3	0.32	0.32	0.09	0.09	0.34	0.26

APPENDIX B

PERTINENT DATA FROM THE WAVE INFORMATION STUDIES PROGRAM

Figures						Page
B 1	Phase	III	wave	rose,	Atlantic City, NJ	B2
B2	Phase	III	wave	rose,	Nagshead, NC	В3
В3	Phase	III	wave	rose,	Daytona Beach, FL	B4
Tables						
B1	Phase	III	Wave	Data,	Atlantic City, NJ:	B2
B2	Phase	III	Wave	Data,	Nagshead, NC	B 3
В3	Phase	III	Wave	Data,	Daytona Beach, FL	B 4
B4	Phase	III	Wave	Data,	Newsport, OR	B5
B 5	Phase	III	Wave	Data,	Half-Moon Bay, CA	B5
В6					ata, Atlantic Coast	В6

Table B1

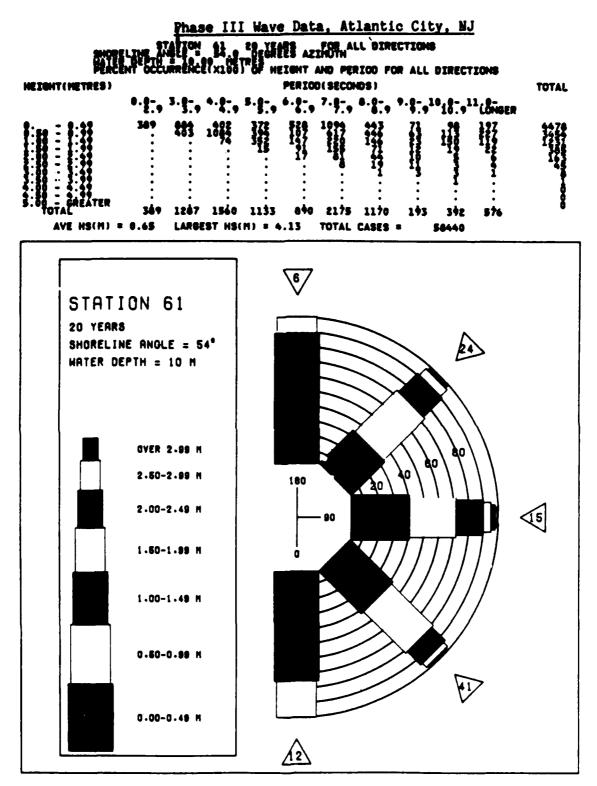


Figure B1. Phase III wave rose, Atlantic City, NJ

Table B2

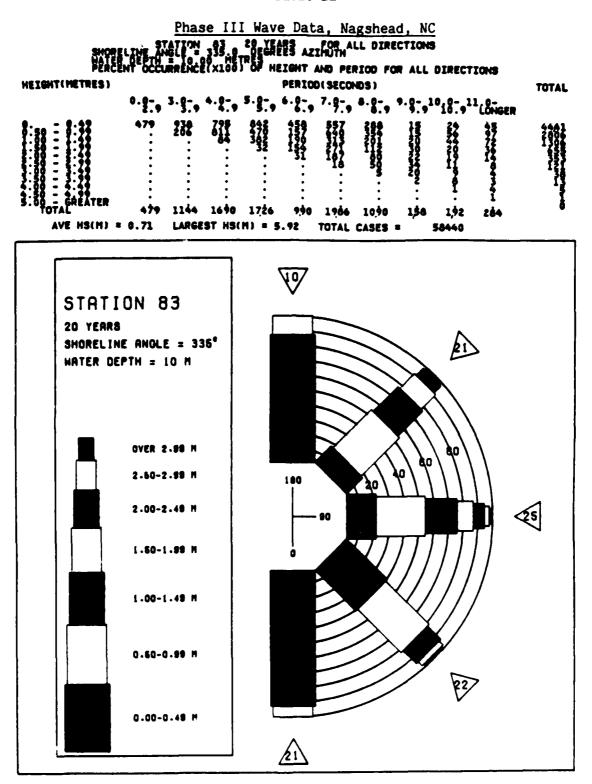


Figure B2. Phase III wave rose, Nagshead, NC

Table B3

Phase III Wave Data, Daytona Beach, FL

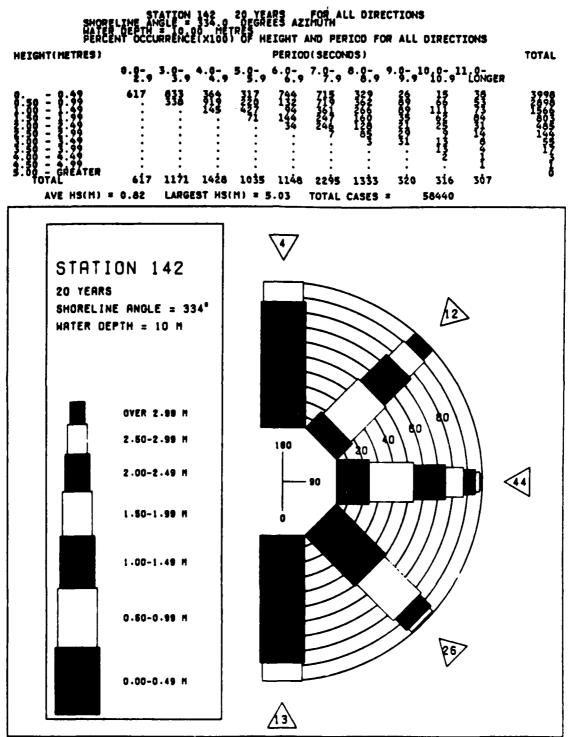


Figure B3. Phase III wave rose, Daytona Beach, FL

Table B4

Phase III Wave Data, Newport, OR

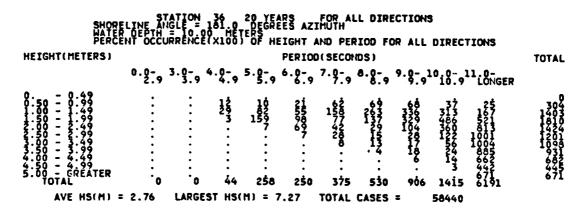


Table B5

Phase III Wave Data, Half-Moon Bay, CA

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2233499 2233499 2333499 2333499	•	:	•	•	**************************************	42 5 :	19 19	1626	138 17 10 7	977 707 465 268	1117 764 495 301
5:00 - GREATER	14	O LARGE	19 St HS(355 M) = 3	509 7 04	686 TOTAL	669 CASES	828	1238 58440	130 5663	130

Table B6

Phase III Duration Data, Atlantic Coast Duration, hr, of Waves with Mg over a Specified Mave Height

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* Duration is shown in hours for M readings "greater than" 0.5, 1.0, etc.

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Table B6 (Concluded)

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APPENDIX C

COMPUTER PROGRAM STRMDIST

	Page
FORTRAN Listing	C2
Sample Output	C14

FORTRAN Listing - Program "STRMDIST"

```
10$$N.J
20$: IDENT: ROCDOPS, OPSMITH
30$: OPTION: FORTRAN
40$:USE:.GTLIT
50$: FORTY
68C
                         HONEYWELL VERSION 1/2/86
70C
            PROGRAM "STRMDIST" READS A WIS PHASE III DATA FILE **
80C
            AND IDENTIFIES STORMS WHERE CONSECUTIVE RECORDS
        **
            HAVE WAVE HEIGHTS EXCEEDING A SPECIFIED THRESHOLD. **
90C
        **
100C
            THE NUMBER, PEAK CONDITIONS AND DURATIONS OF THESE **
            STORMS ARE THEN TABULATED. STORMS ONLY & HOURS
118C
        **
            APART ARE CONSIDERED AS A SINGLE EVENT. THE
120C
138C
            PROGRAM ALSO FITS AN EXTREMAL TYPE I AND A WEIBULL **
140C
            DISTRIBUTION TO THE PEAK WAVE HEIGHTS AND THE DUR- **
        * *
            ATIONS AND REPORTS THE PARAMETERS OF EACH.
150C
160C
         DIMENSION DUR(999), HPEAK(999), TPEAK(999), DPEAK(999), DTPEAK(999)
178
         INTEGER DATM, HSEA, TSEA, DSEA, HSWL, TSWL, DSWL, STNO, DATIM, H, T, D
180
         INTEGER STMNO.DN.HPK.TPK.DPK.HPEAK.TPEAK.DPEAK.DUR.H1.H2.RECNO
190
         INTEGER DTPEAK, DATIME, DTPK, FLAG, NOREC, NOYRS, TMIN, TMAX
200
         INTEGER YR, YRP, MO, MOP, DY, DYP, TM, TMP, YRDIFF, MODIFF, DYDIFF,
210
228
        &TMDIFF, HPKMIN, HPKMAX
230
         CHARACTER*64 FNAME
240
         CHARACTER*8 VARIABLE
250
         CHARACTER*8 VARIABLE
260C
270C
        ** READ WIS DATA FILE AND WRITE FILE OF STORMS **
        ** EXCEEDING 1ST WAVE HEIGHT THRESHOLD, H1 (CM) **
288C
290C
300
         FNAME = "NAGSHEAD, NORTH CAROLINA"
310
         NOREC=58440
         NOYRS=20
320
330
         H1 = 300
340C
            K = THE STORM NO. ASSIGNED TO CONSECUTIVE RECORDS
350
         K=0
        ** J = THE NO. RECORDS WHERE H > H1
360C
          J = \emptyset
370
380
         YRP=999
390
         MOP=999
          DYP=999
400
410
          TMF=999
          CALL ATTACH(01,"/A3083;",1,0,1STAT)
428
          ISTAT=FLD(6,6,ISTAT)
430
          IF (ISTAT.NE.0) GO TO 900
448
          CALL FMEDIA(07,6)
450
455
          CALL FMEDIA (08,6)
460
          CALL FMEDIA (09,6)
479
          READ(1,10) YR, MO, DY, TM, HSEA, TSEA, DSEA, HSWL, TSWL, DSWL
480
       10 FORMAT (2X,412,616)
498C
```

を利用してこのこののこの間である。 (A provided 12 provide

```
** COMPUTE COMPOSITE SEA AND SWELL WAVE HEIGHT, 1ST RECORD **
599C
510C
         H=INT(SORT(FLOAT(HSEA) **2+FLOAT(HSWL) **2))
528
         IF(H.LT.H1) 80 TO 30
539
548
         J=1
558
         K=1
568C
            SET COMPOSITE PERIOD AND DIRECTION TO THAT OF SEA OR
578C
                                                                         **
            SWELL, WHICHEVER HAS A HIGHER INCREMENTAL WAVE HEIGHT
580C
598C
         IF (HSEA.GT. HSWL) 60 TO 15
600
618
         T=TSWL
         D=DSWL
620
         60 TO 20
638
      15 T=TSEA
648
         D=DSEA
650
      28 WRITE(7,25) K, YR, MO, DY, TM, H, T, D
660
      25 FORMAT (2X, 14, 1X, 412, 316)
670
688
         YRP=YR
698
         MOP=MO
788
         DYP=DY
710
         TMP=TM
728
      30 READ(1,10,END=200) YR,MO,DY,TM,HSEA,TSEA,DSEA,HSWL,TSWL,DSWL
730C
        ** COMPUTE COMPOSITE SEA AND SWELL WAVE HEIGHT
740C
750C
         H=INT(SQRT(FLOAT(HSEA) **2+FLOAT(HSWL) **2))
760
      32 IF(H.LT.H1) GO TO 190
778
780
         J=J+1
790C
800C
        ** SET COMPOSITE PERIOD AND DIRECTION TO THAT OF SEA OR
810C
        ** SWELL. WHICHEVER HAS A HIGHER INCREMENTAL WAVE HEIGHT
820C
830
         IF (HSEA.GT. HSWL) GO TO 35
         T=TSWL
848
850
         D=DSWL
860
         60 TO 40
870
      35 T=TSEA
889
         D=DSEA
898
      40 YRDIFF=YR-YRP
900
         MODIFF=MOP-MO
910
         DYDIFF=DYP-DY
         TMDIFF=TMP-TM
920
930C
           *** CHECK FOR CONSECUTIVE RECORDS (SAME STORM) ***
940C
950C
960C
                CONSECUTIVE RECORDS, SAME DAY
         IF (YRDIFF.EQ. 0. AND. MODIFF.EQ. 0. AND. DYDIFF.EQ. 0. AND.
970
        &TMDIFF.EQ.-3) 60 TO 45
980
990C
           ** CONSECUTIVE RECORDS, DAY END
          IF (YRDIFF.EQ. 0. AND. MODIFF.EQ. 0. AND. DYDIFF.EQ. -1. AND.
1000
         &TMDIFF.EQ.21) 60 TO 45
1010
                 CONSECUTIVE RECORDS, MONTH END
1020C
```

THE STATE OF THE S

```
IF (YRDIFF.EQ. 8. AND. MODIFF. EQ. -1. AND. DYDIFF. EQ. 27. AND.
1030
1848
          &TMD1FF.EQ. 21) 60 TO 45
           IF (YRDIFF.EQ. @. AND. HODIFF.EQ. -1. AND. DYDIFF.EQ. 28. AND.
1050
1868
          &TMDIFF.EQ.21) 60 TO 45
           IF (YRDIFF. EQ. Q. AND. MODIFF. EQ. -1. AND. DYDIFF. EQ. 29. AND.
1979
1080
          &TMDIFF.EQ.21) 60 TO 45
           IF (WRDIFF.EQ. 8. AND. MODIFF.EQ. -1. AND. DYDIFF.EQ. 38. AND.
1898
          &TMDTFF.EG.21) 60 TO 45
1100
                 CONSECUTIVE RECORDS, YEAR END
1110C
           IF (YRDIFF.EQ. 1. AND. MODIFF. EQ. 11. AND. DYDIFF. EQ. 36. AND.
1120
1138
          &TMD1FF.EQ. 21) GO TO 45
114BC
                 CHECK FOR RECORDS 6 HRS APART AND ADJUST RECORD
1158C
                 BETWEEN SUCH THAT THE PROGRAM SEES ONE CONT-
            ***
1168C
                                                                       ...
                 INUOUS STORM (IGNORING THE ONE RECORD BELOW
1178C
            ...
                                                                       ...
1188C
            ***
                 THE THRESHOLD)
1198C
                  RECORDS 6 HRS APART, SAME DAY
1208C
           IF (YRDIFF.EQ. @. AND. MODIFF.EQ. @. AND. DYDIFF.EQ. @. AND.
1219
          &TMDIFF.EQ.-6) GO TO 47
1220
             ** RECORDS & HRS APART, DAY END
1238C
1248
           IF (YRDIFF.EQ. 8. AND. HODIFF.EQ. 8. AND. DYDIFF.EQ. -1. AND.
1250
          &TMD[FF.EQ.18) 60 TO 47
                 RECORDS 6 HRS APART, MONTH END
126BC
                                                        ..
           IF (YRDIFF.EQ. 0. AND. HODIFF.EQ. -1. AND. DYDIFF.EQ. 27. AND.
1276
1280
          &TMDIFF.EQ.18) 60 TO 47
           IF (YRDIFF.EQ. 0. AND. MODIFF.EQ. -1. AND. DYDIFF.EQ. 20. AND.
1298
         &TMDIFF.EQ.19) 80 TO 47
1300
          IF (YRDIFF.EQ. 0. AND. MODIFF.EQ. -1. AND. DYDIFF.EQ. 29. AND.
1310
         &TMDIFF.EQ.18) GO TO 47
1328
           IF (YRDIFF. EQ. 0. AND. MODIFF. EQ. -1. AND. DYDIFF. EQ. 30. AND.
1330
          &TMDIFF.EQ.18) 60 TO 47
1340
                 RECORDS & HRS APART, YEAR END
1358C
                                                        ..
           IF (YRDIFF.EQ. 1. AND. MODIFF.EQ. 11. AND. DYDIFF.EQ. 30. AND.
1360
          &TMDIFF.EQ.18) 60 TO 47
1370
1380
           K=K+1
           60 TO 45
1390
       47 BACKSPACE 1
1480
1410
           BACKSPACE 1
           READ(1,18) YR, MO, DY, TM, HSEA, TSEA, DSEA, HSWL, TSWL, DSWL
1428
1438
           H=H1
           60 TO 32
1440
        45 WRITE(7,25) K,YR,MO,DY,TM,H,T,D
1450
           YRP=YR
1460
1470
           MOP-MO
           DYP-DY
1488
           TMP=TM
1498
1500 190 60 TO 30
1518C
         ** FILE CODE 7 INCLUDES RECORDS WHERE H IS GREATER THAN
1528C
            THE FIRST WAVE HEIGHT THRESHOLD HI. CONSECUTIVE
1538C
            RECORDS SHARE A COMMON "STORM NUMBER", K.
1548C
1550C
```

```
1560C
1570C
         ** READ FILE OF STORMS , COMPUTE DURATIONS AND
1588C
              IDENTIFY PEAK CONDITIONS
1598C
1688 288 RECNO=J
1619
           REWIND 7
1628
           DN=8
           HPK=8
1630
1648
           J=1
1658
        95 READ (7,97,END=400) STNO, DATM, H, T, D
1669
        97 FORMAT(2X,14,1X,18,316)
1670
           IF(STNO.EQ.J) 60 TO 350
1688
           DUR(J)=DN
1698
           DTPEAK(J) = DTPK
1788
           HPEAK (J) =HPK
1719
           TPEAK(J)=TPK
1728
           DPEAK (J) = DPK
1738
           J=J+1
1748
           DN=8
1750
           HPK=8
     358 DN=DN+3
1769
1778
           IF(H.LE.HPK) GO TO 95
1780
           DTPK=DATM
1798
           HPK≈H
1800
           TPK=T
1810
           DPK=D
1820
           GO TO 95
1836
     400 DUR(J)=DN
1846
           DTPEAK(J) = DTPK
1850
          HPEAK (J) = HPK
           TPEAK(J)=TPK
1868
1879
          DPEAK (J) = DPK
1886
           POISSON=FLOAT(J)/FLOAT(NOYRS)
1898
          PERCENT=RECNO+100./NOREC
1988
           TMIN=DUR(1)
          HPKMIN=HPEAK(1)
1918
1928
          HPKMAX=H1
1938
          TMAX=3
1948
          TSUM=8.8
1958
          HPKSUM=0.0
1960
          DO 700 I=1,J
1970
          TSUM=TSUM+FLOAT(DUR(I))
1980
          HPKSUM=HPKSUM+FLOAT (HPEAK(I))
          IF (HPEAK(I).LT.HPKMIN) HPKMIN≈HPEAK(I)
1998
2000
           IF (HPEAK(I).GT.HPKMAX) HPKMAX=HPEAK(I)
2916
          IF (DUR(I).LT.TMIN) TMIN=DUR(I)
2020
          IF(DUR(I).GT.TMAX) TMAX=DUR(I)
      708 CONTINUE
2030
2040
          HPKMEAN=HPKSUM/FLOAT(J)
2950
          TMEAN=TSUM/FLOAT(J)
2060
          TDIFFSUM=0.0
2070
          HDIFSUM=0.0
2088
          DO 718 I=1,J
```

```
2898
          HDIFSQ=(FLOAT(HPEAK(I))-HPKMEAN)**2
2188
          HDIFSUM=HDIFSUM+HDIFSQ
          TDIFSQ=(FLOAT(DUR(I))-TMEAN)++2
2110
2120
          TDIFFSUM=TDIFFSUM+TDIFSQ
2138 718 CONTINUE
          STDEVT=SQRT(TDIFFSUM/FLOAT(J-1))
2140
          STDEVH=SQRT(HDIFSUM/FLOAT(J-1))
2150
2168C
              PRINT TABLE OF STORM PARAMETERS
2178C
2180C
2190
          WRITE(6,448) FNAME
      448 FORMAT(1H1,///,25x, "ANALYSIS OF STORM DURATION",
2288
         1//,8X,"DATA FILE: ",A64)
2210
2228
          WRITE(6,450) H1
2238
      450 FORMAT(//,1X, "STORM NO.",2X, "DATE/TIME OF PEAK ",2X,
         & DURATION H> , 13,2x, "PEAK H",2x, "PEAK T",2x, "PEAK DIR",/)
2240
          DO 586 L=1,J
2258
          WRITE(6,478) L, DTPEAK(L), DUR(L), HPEAK(L), TPEAK(L), DPEAK(L)
2268
2279
          WRITE(9,470) L,DTPEAK(L),DUR(L),HPEAK(L),TPEAK(L),DPEAK(L)
2288 478 FORMAT(4X,13,18X,18,8X,14,12X,13,6X,12,6X,13)
2290
     500 CONTINUE
2388
          WRITE (6,518) PDISSON
2310 510 FORMAT(/,4x,F5.2," STORMS PER YEAR")
          WRITE(6,28) H1, RECNO, PERCENT, NOREC
2328
2338
       28 FORMAT(/,4x,"NO. RECORDS WHERE H > ",13," = ",15.
         &" (",F4.1,"% OF ",I5," RECORDS)")
2349
2358
          WRITE(6,719) HPKMIN, HPKMAX, HPKMEAN, STDEVH
2368
      719 FORMAT(/,4x,"MIN. PEAK H = ",13," MAX. = ",14,
              MEAN = ",F5.1," STD. DEV. = ",F5.1)
2370
          WRITE (6,720) THIN, THAX, THEAN, STDEVT
2380
2398
      728 FORMAT(/,4x,"MIN. DURATION = ",13,"
              MEAN = ",F5.1," STD. DEV. = ",F5.1)
2488
2418
          WRITE (6,475)
2428 475 FORMAT(//,4X,"THE DATE/TIME IS YRMODYHR, DURATION IS IN HOURS,
         &H (HEIGHT) IS IN CM,",/,4X,"T (PERIOD) IS IN SEC AND DIRECTION",
2438
2448
         &" IS IN DEGREES RELATIVE TO THE SHORELINE")
          VARIABLE= 'PEAK H
2458
          CALL PROBLIST (VARIABLE, J, HPEAK, POISSON)
246
2478
          VARIABLE='DURATION'
2480
          CALL PROBDIST (VARIABLE, J. DUR, POISSON)
2498
          60 TO 620
2500 900 PRINT 901
2510 901 FORMAT(1x,19HATTACH UNSUCCESSFUL)
2528
          CALL DETACH(01,,)
2538 628 STOP
2540
          END
2558C
2568C
2570C
          SUBROUTINE PROBDIST (VARIABLE, N. HS, LAMBDA)
2588
2598C
       SUBROUTINE PROBDIST ADAPTED 1/86 BY ORSON P. SMITH FROM
2600C
       PROGRAM "WAVDISTI". 11/85 VERSION BY ROBERT B. LUND
2618C
       DESIGN BRANCH-COASTAL ENGINEERING RESEARCH CENTER
```

```
2620C U.S. ARMY ENGINEERS WATERWAYS EXPERIMENT STATION
2630C P.O. BOX 631
2648C VICKSBURG, MS 39188-8631
2658C FOR FURTHER INFORMATION CONCERNING THE APPLICATION
2668C OF "WAVDIST1", CALL....
2670C ROBERT B. LUND (601)-634-2068 FTS: 2068
2688C ORSON P. SMITH (681)-634-2813 FTS:542-2813
2698C DOYLE L. JONES (681)-634-2869 FTS:542-2869
2788C
2718C FORTRAN 4
                      HONEYWELL DPS-8
2720C REF: "RELIABILITY OF LONG-TERM WAVE CONDITIONS PREDICTED WITH DATA SETS
            OF SHORT DURATION" CETN-1-5
2748C REF: "HANDBOOK OF MATHEMATICAL FUNCTIONS" BY ABRAHOWITZ AND SEGUN
2750C REF: "EXTREMAL PREDICTION IN WAVE CLIMATOLOGY" BY BORGMAN AND RESID
2768C REF. "LONG-TERM DISTRIBUTIONS OF OCEAN WAVES."
2778C
            ISAACSON AND MACKENZIE
2788C
           N = NUMBER OF STORMS
2798C
         RET = RETURN PERIOD
2800C
2810C LAMBDA = POISSON LAMBDA PARAMETER (AVERAGE NO. STORMS PER YEAR)
2828C
          HS = THE INDEPENDENT VARIABLE
        DIFF = THE RESIDUAL FOR EACH DATA POINT
2838C
       YACT = THE PROBABILITY AS ESTIMATED BY THE PLOTTING FORMULA M/K+1
2840C
        VEST = THE PROBABILITY AS ESTIMATED BY THE DISTRIBUTION
2850C
2868C ALPHA = THE ARRAY OF LOCATION PARAMETERS FOR THE DISTRIBUTIONS
        BETA - THE ARRAY OF SCALE PARAMETERS FOR THE DISTRIBUTIONS
2878C
           A = THE SLOPE OF EACH "PLOTTED LINE"
2880C
           B = THE Y-INTERCEPT OF EACH "PLOTTED LINE"
2898C
           C = THE ARRAY OF COEFFICIENTS FOR THE GAMMA INTEGRAL EXPANSION
2968C
          ST = THE SUM OF THE SQUARE RESIDUALS
2918C
        CORR = THE NON-LINEAR CORRELATION FOR EACH DISTRIBUTION
2928C
         STE = THE STANDARD ERROR OF THE ESTIMATE OF Y ON X
2938C
294BC
         MSD = THE MEAN SQUARE DEVIATION
2958
2960C DECLARATION OF VARIABLES, FUNCTIONS, AND CHARACTERS
            DIMENSION YACT (999,3), YEST (999,3), DUM1 (999), DUM2 (999), HS (999)
2970
2988
            DIMENSION YAVG(3), CORR(3), ALPHA(4), BETA(4), VAR(4), DM(3)
2990
            DIMENSION RET(5), CHS(5,3), A(3), B(3), ST(3), SB(3), STE(3)
3000
            DIMENSION STDEV(3)
            REAL MEAN(3), MSD(3)
3010
3020
            REAL LAMBDA
3030
          INTEGER HS
3848
            F1(X) = EXP(-EXP(-(X-EPSI)/PHI))
3050
3866
            F2(X)=1.0-EXP((-(X/SIGMA)**C))
3070
            F3(X) = EXP(-((SIGMA2/X) ++U))
3088
            CHARACTER*28 IFLAG(4)
3898
3100
            CHARACTER+17 DEF
            CHARACTER#34 FORM(3)
3110
            CHARACTER+24 TITLE
3120
3130
            CHARACTER*1 LOGIC
3148
            CHARACTER+60 BOX(16)
```

```
3150
            CHARACTER*B VARIABLE
3168C INITIALIZATION OF STRINGS AND CONSTANTS
            IFLAG(1) = 'EXTREMAL TYPE I'
3170
3180
            IFLAB(2) = 'WEIBULL
3198
            IFLAG(3) = 'LOS EXTREMAL'
3200
            DEF='F(x)=Pr(X(x)=
3210
            FORM(1) = 'EXP(-EXP(-(x-EPSI)/PHI))'
            FORM(2) = '1 - EXP(-(x/BETA) + ALPHA)'
3228
3230
            FORM(3) = 'EXP(-(BETA/x) + ALPHA)'
            TITLE='LEAST SQUARES RESULTS - '
3240
3250
3260
            DATA RET /5.0,10.0,25.0,50.0,100.0/
3278
            EULER=.5772156649
3288
            C2=.7796968
3298
3388C
         ** SET LOGIC = 'Y' FOR PRINTOUT OF RESIDUAL TABLES
3310C
3320C
          LOGIC='N'
3338
3340C
       RANK DATA AND ASSIGN A PROB. OF NON-EXCEEDENCE TO EACH
3350
            CALL ORDER (HS.N)
3360
            DO 25 I=1,N
3378
            DO 25 K=1,3
3380
            YACT(I,K)=FLOAT(I)/FLOAT(N+1)
3398
       25 CONTINUE
3488
3410C INITIALIZE VARIABLES FOR LEAST SQUARES FIT OF THE DISTRIBUTIONS
3428
            SX=0
3439
            SY=0
3440
            SXX=0
3450
            SLX-8
3468
            SLLY=8
3478
            SLXX=8
3480
            SLLQY=0
3498
            SXLLY=0
3500
            SLXLLY = 8
3518
            TOOBIG=8
3520
3530C CALCULATE SUMS FOR THE LEAST SQUARES METHOD
3540
            DO 40 J=1.N
3550
            SX=SX+HS(J)
3540
            SY=SY+YACT(J,1)
            SXX=SXX+HS(J)++2
3570
3588
            SLX=SLX+ALOG(HS(J))
3598
            SLXX=SLXX+(ALOG(HS(J))) ##2
3600
            SLLY=SLLY-ALOG(-ALOG(YACT(J,1)))
3618
            SLLQY=SLLQY+ALOG(-ALOG(1.8-YACT(J,1)))
3629
            SXLLY=SXLLY-HS(J) *ALOG(~ALOG(YACT(J,1)))
3630
            SLXLLY=SLXLLY-ALOG(HS(J)) *ALOG(-ALOG(YACT(J,1)))
3648 48
            TOOBIG=TOOBIG+ALOG(HS(J)) * (ALOG(-ALOG(1.8-YACT(J,1))))
3650
3660C CALCULATE SLOPE AND INTERCEPT OF EACH "PLOTTED LINE"
3670
            A(1) = (N+SXLLY-SX+SLLY)/(N+SXX-SX++2)
```

```
3988
             A(2) = (N*TOOBI6-SLX*SLLQY) / (N*SLXX-SLX+*2)
3690C
             A(3) = (N+SLXLLY-SLX+SLLY)/(N+SLXX-SLX++2)
3788
             B(1) = (SXX + SLLY - SXLLY + SX) / (N + SXX - SX + + 2)
             B(2) = (SLXX + SLLQY - TOOBIG + SLX) / (N + SLXX - SLX + + 2)
3710
372BC
             B(3) = (SLXX+SLLY-SLXLLY+SLX)/(N+SLXX-SLX++2)
3730C CALCULATE PARAMETERS OF EACH DISTRIBUTION FROM SLOPE AND INTERCEPT DATA
3748
             PHI=1.0/A(1)
3758
             EPSI=-B(1)/A(1)
3769
             C=A(2)
3778
             SIGMA=EXP(-B(2)/A(2))
3788C
             U=A(3)
3790C
             SIGMA2=EXP(-B(3)/A(3))
3888
3810C ASSIGN ARRAYS ALPHA AND BETA THE PARAMETERS OF EACH DISTRIBUTION
3828C FOR EASY PRINTOUT OF DATA
3838
             ALPHA(1) = EPSI
3840
             BETA(1)=PHI
3858
             ALPHA(2)=C
3868
             BETA(2)=SIGMA
3870C
             ALPHA(3)=U
3880C
             BETA(3)=SIGMA2
3890C CALCULATE PROBABILITY AS ESTIMATED BY DISTRIBUTION
3900
             DO 100 J=1,N
3918
             YEST(J,1)=F1(HS(J))
3928
             YEST(J,2)=F2(HS(J))
3938C
             YEST(J,3) = F3(HS(J))
3940 100 CONTINUE
3950
3968C CALCULATE AVERAGE PROBABILITY AND CORRELATION COEFFICIENTS
3978
             DD 110 K=1,2
3988
             YAVG(K) = SY/FLOAT(N)
3990
             MSD (F) = 0
4000
             ST(K)=0
4810 110
             SB(K)=0
4020
4030
            DO 120 k=1,2
4848
             DO 130 I=1.N
4858
             ST(K) = ST(K) + (YACT(I,K) - YEST(I,K)) **2
4868 138
             SB(K) = SB(K) + (YACT(I,K) - YAV6(K)) + *2
4678
             IF( (1.0-ST(k)/SB(K)) .LT. 0) CORR(K)=0.
             IF( (1.0-ST(K)/SB(F)) .LT. 0) GO TO 125
4688
4098
             CORR(K) = SQRT(1.0-ST(K)/SB(k))
             IF( N .EQ. 2) 60 TO 120
4188 125
            STE(K) = SQRT(ST(Y) / (N-2))
4118
4120 120
            CONTINUE
4130
4140C CALCULATE DATA FOR RETURN PERIOD TABLES
4158
             DO 57 J=1,5
            FROB=1.0-1.3/(LAMBDA*RET(J))
4160
4170
             IF (PROF .LE. 0) PROB=.0000001
            CHS(J,1) =-ALOG(-ALOG(PROB)) +PHI+EPSI
4198
4198
            CHS(J,2) = (-ALOG(1.0-PROB)) **(1.0/C) *SIGMA
             CHS(J.3)=S15M42/((-ALOG(PROB))++(1.0/U))
4288C
```

```
4218
       57 CONTINUE
4228
4230C CALCULATE MEAN SQUARE DEVIATION FOR EACH DISTRIBUTION
4248
             DO 58 I=1.N
4258
             21=YACT(1,1)
4268
             Z2=EPSI-(ALOG(-ALOG(Z1)))*PHI
4278
             I3=BETA(2) + ((-ALOS(1-I1)) + + (1.0/ALPHA(2)))
42BBC
             Z4=BETA(3)/((-ALO6(Z1))++(1.0/ALPHA(3)))
4298
             MSD(1) = MSD(1) + (22 - HS(1)) + *2
4300
             MSD(2) = MSD(2) + (Z3 - HS(I)) **2
4310C
             MSD(3) = MSD(3) + (24 - HS(1)) + +2
4320
      58 CONTINUE
4338
             MSD(1) = MSD(1) / (N + PHI + + 2)
4348
             MSD(2) = MSD(2) / (N*BETA(2) **2)
4358C
             MSD(3) = MSD(3) / (N*BETA(3) **2)
4368
4370C CALCULATE MEAN AND VARIANCE FOR EACH DISTRIBUTION
4388
             MEAN(1) = EPSI+EULER * PHI
4398
             VAR(1)=1.6449341*PHI**2
             PARA=1.8+1.8/C
4488
             CALL GAMMA (PARA, WME)
4418
4420
             MEAN(2)=SIGMA+WME
4438
             FAC1=SIGMA++2*WME++2
4449
             PARA=1.0+2.0/C
4458
             CALL BAMMA (PARA, WV2)
4468
             FAC2=SIGMA++2+WV2
4478
             VAR(2)=FAC2-FAC1
4488C
             PARA=1.0-1.0/U
4498C
             CALL GAMMA (PARA, HPC)
4588C
             MEAN(3)=SIGMA2+HPC
             PARA=1.8-2.8/U
4518C
             CALL GAMMA (PARA, HPD)
4529C
             VAR(3)=SI6MA2++2+HPD-MEAN(3)++2
4538C
4548
4550C WRITE OUT THE DATA FOR EACH DISTRIBUTION
4568
           WRITE (6, 136)
4578 136 FORMAT (1H1)
             WRITE(6,135) TITLE, VARIABLE
4588
4598 135
             FDRMAT(///,16%,A26,A8,///)
4688
             DO 150 K=1,2
           STDEV(K) = SQRT(VAR(K))
4610
4620
             WRITE (6, 160) IFLAG (K), DEF, FORM (K)
4638 168
             FORMAT (15X, A30, //, 1X, A17, 2X, A34)
             IF( F .EQ. 1) WRITE(6,159) EPSI, PHI
4649
             FORMAT(1X, "EPSI="6X, F10.3, /, 1X, "PHI=", 7X, F10.3)
4550 159
             IF( K .GT. 1) WRITE(6,161) ALPHA(K), BETA(K)
4668
4678 161
             FORMAT(1x, "ALPHA=", 6x, F18.3, /, 1x, "BETA=", 6x, F18.3)
4688
             WRITE(6,162) MEAN(K), VAR(K), STDEV(K)
4698 162
             FORMAT(1X, "MEAN=", 6X, F18.3, /, 1X, "VARIANCE=", 2X, F18.3,
          \frac{1}{4}, 1x, "STD. DEV. = ",2x,F7.3)
4788
             IF ( LOGIC .EQ. 'N') GO TO 171
4718
4720
             DO 178 I=1,N
4730
             DUM1(I)=YACT(I,K)
```

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```
4748 178
             DUM2(I)=YEST(I,K)
4758
             L2=N
4760
             CALL RESIDUAL (HS, DUM1, DUM2, L2)
4770 171.
             WRITE(6,163) CORR(K),ST(K)
             FORMAT(/,1x, "NON-LINEAR CORRELATION IS",5x,F19.7,/
4788 163
             ,1x, "SUM SQUARE RESIDUALS IS",6x,F11.7)
4798
             IF( N .EQ. 2 ) 8D TO 167
4686
             WRITE(6,164) STE(K)
4810
4828 164
             FORMAT(1x, "STANDARD ERROR IS", 13x, F10.7)
4838 167
             WRITE(6,166) MSD(K)
4848 166
             FORMAT(1X, "MEAN SQUARE DEVIATION IS", 6X, F10.7, ///)
4859 207
             WRITE(6,208) VARIABLE
4860 208
             FORMAT(7X, "RETURN PERIOD TABLE", /, 6X, "YEAR", 13X, AB)
             DO 211 J=1,5
4878
             WRITE (6,212) RET (J), CHS (J,K)
4888
4898 212
             FORMAT(1X,F9.2,8X,F9.2)
4988 211
             CONTINUE
4918
             WRITE(6,165)
4928 165
             FORMAT (////)
4938 158
             CONTINUE
4948
             RETURN
             END
4950
4968
4978
4988
4998
5000C SUBROUTINE TO PUT NUMBERS IN ORDER BY ASCENDING X
             SUBROUTINE ORDER (X,N)
5010
             DIMENSION X(N)
5020
5030
           INTEGER X,TX
             DD 28 K=2,N
5948
5050
             J=N-K+2
             DO 10 I=1,J-1
5968
             IF( X(I) .LT. X(I+1)) GO TO 18
5070
5888
             TX = X(I)
             X(I)=X(I+1)
5898
5188
             X(I+1)=TX
             CONTINUE
5110 10
5128 28
             CONTINUE
5130
             RETURN
             END
5148
5150
5160
5178
5180C SUBROUTINE TO HELP PRINT CUT DATA
5190
             SUBROUTINE RESIDUAL (X, YACT, YEST, N)
             DIMENSION X(N), YACT(N), YEST(N), DIFF(200)
5200
           INTEGER X
5210
             55R=0
5228
             DO 18 1=1,N
5238
             DIFF(I) = (YACT(I) - YEST(I)) **2
5240
5250 10
             SSR=SSR+DIFF(I)
             WRITE (6,15)
5268
```

```
FORMAT (//, 1X, "
                               XVALUE
                                                                    DIFF ",/,)
5278 15
                                            YVALUE
                                                         YEST
5280
            DO 25 I=1.N
5298
            WRITE(6,20) X(I), YACT(I), YEST(I), SQRT(DIFF(I))
5300
          WRITE(8,28) X(I), YACT(I), YEST(I), SQRT(DIFF(I))
5318 28
            FORMAT(1X, 111, F11.4, F11.4, F11.4,/,)
5320 25
            CONTINUE
5330
            RETURN
5340
            END
5350
5368
       SUBROUTINE TO EVALUATE THE GAMMA FUNCTION
5370C
       PROGRAM ADJUSTS ALPHA TO BE BETWEEN 1.0 AND 2.0
5380C
539BC
       AND THEN MULTIPLIES BY GF TO COMPENSATE
5400
            SUBROUTINE GAMMA (ALPHA, AREA)
            DOUBLE PRECISION C(25), SUM
5418
5420
            BF=1.0
5438
            IF (ALPHA) 1,2,3
5448
5458 2
            PRINT, 'TROUBLE IN SAMMA'
5468
            AREA=1.8
5478
            60 TO 200
5480
5490C FOR BAMMA OF A POSITIVE NUMBER
            M=INT(ALPHA)
5500 3
            EPSI=ALPHA-FLDAT(M)
5518
5520
            IF( M .EQ. 0) GF=GF/ALPHA
5530
            IF ( M .EQ. 8) ALPHA=ALPHA+1.8
5548
            IF( M .EQ. 0) GO TO 100
5559
            IF ( M .EQ. 1) 6F=1.0
5568
            IF( M .EQ. 1) GO TO 100
5570
            DO 10 I=2.M
            GF=GF*(FLOAT(I-1)+EPSI)
5588 18
559B
            ALPHA=1.0+EPSI
            60 TO 190
5600
5610
5620C FOR GAMMA OF A NEGATIVE NUMBER
5630 1
            M=INT (ALPHA)
5649
            EPSI=ALPHA-FLOAT(M)
5650
            DO 20 I=1,2~M
5668
            J=H+(I-1)
5670 20
            GF=GF/(EPSI+FLOAT(J))
            ALPHA=EPSI+2.0
5680
5698
       COEFFICIENTS FOR SERIES EXPANSION OF THE GAMMA INTEGRAL
5700C
5710C SEE HANDBOOK OF MATHEMATICAL FUNCTIONS BY ABRAMOWITZ AND SEGUN
5720 180
            5730
            C(2) = .5772156649015329
            C(3) = -.6558780715202538
5748
5750
            C(4) = -.0420026350340952
5769
            C(5) = .1665386113822915
5770
            C(6) = -.0421977345555443
5788
            C(7) = -.889621971527887
5798
            C(8) = .887218943246663
```

```
5888
             C(9) = -.8811651675918591
             C(18) = -. 8882152416741149
5816
5820
             C(11)=.0001280502823882
             C(12)=-.0000201348547807
5839
5840
             C(13) = -.0000012504934821
             C(14) = .0000011330272320
5850
             C(15) = -.0888882856338417
5868
5878
             C(16)=6.116895E-89
5888
             C(17)=5.8928875E-89
             C(18)=-1.1812746E-09
5898
5988
             C(19)=1.943427E-10
5918
             C(20) = 7.7823E - 12
             C(21) = -3.69680E-12
5928
             C(22) = 5.1E-13
5930
             C(23) = -2.86E - 14
5948
             C(24) = -5.4E - 15
5950
             C(25)=1.4E-15
5968
5978
      SUM SERIES
5980C
5998
             SUM=0.0
             DO 50 K=1,25
6000
             SUM=SUM+C(K) * (ALPHA++K)
6010
6020 50
             AREA=8F/SUM
             RETURN
6939 200
6848
             END
6050$: EXECUTE
6060$:LIMITS:30,100K
6070$: FILE: 07, X7R, 5L, NEW, STRMFILE
69884: FILE: 00, XOR, 5L, NEW, DISTFILE
6090$:FILE:09, X9R, 5L, NEW, A83DST30
6188$: ENDJD8
```

ANALYSIS OF STORM DURATION

DATA FILE

NAGSHEAD, NORTH CAROLINA

STORM NO.	DATE/TIME OF PEAK	DURATION H>350	PEAK H	PEAK T	PEAK DIR
1	56011015	27	449	11	98
	56092718	6	379	18	101
2 3 4	56102703	3	461	11	83
	56102808	9	495	11	75
5	56103100	3	377	8	69
6	56103109	6	364	11	74
7	58182183	3	352	10	91
8	58102118	9	488	10	75
9	58102212	27	508	11	88
10	60020100	9	386	10	93
11	60020118	3	351	10	101
12	60103121	3	354	18	97
13	61102421	6	387	11	99
14	62030721	18	459	11	114
15	62030909	30	591	13	94
16	621128 8 6	84	466	12	97
17	62120200	3	351	9	88
18	62120212	15	371	- 9	68
19	63020421	6	363	9	88
20	64092221	21	391	10	111
21	66061306	3	403	10	105
22	68011121	12	488	10	100
23	68022512	15	385	10	116
24	69022109	3	355	10	114
25	69939306	3	360	11	101
26	7010271B	12	364	10	193
27	72052700	15	365	10	94
28	73021112	33	465	11	111
29	73021300	3	372	10	101
30	73022806	9	379	10	107
31	73120906	3	352	9	66
32	75012118	6	389	10	106
33	75070106	9	399	10	112
34	75070215	9	438	11	104
35	75112415	9	375	10	107
36	75112506	3	397	10	102

^{1.80} STORMS PER YEAR

NO. RECORDS WHERE H > 350 = 149 (0.3% OF 58440 RECORDS)

MIN. PEAK H = 351 MAX. = 591 MEAN = 401.2 STD. DEV. = 53.4

MIN. DURATION = 3 MAX. = 84 MEAN = 12.2 STD. DEV. = 14.9

THE DATE/TIME IS YRMODYHR, DURATION IS IN HOURS, H (HEIGHT) IS IN CM, T (PERIOD) IS IN SEC AND DIRECTION IS IN DEGREES RELATIVE TO THE SHORELINE

LEAST SQUARES RESULTS - DURATION

EXTREMAL TYPE I

F(X)=PR(X <x)=< th=""><th>EXP(-EXP(-(X-EPSI)/PHI))</th></x)=<>	EXP(-EXP(-(X-EPSI)/PHI))
EPSI=	3.918
PHI=	15.246
MEAN=	12.718
VARIANCE=	382.333
STD. DEV. =	19.553

NON-LINEAR CORRELATION IS 8.8720603
SUM SQUARE RESIDUALS IS 8.6796928
STANDARD ERROR IS 0.1413894
MEAN SQUARE DEVIATION IS 0.3527218

RETURN PERIOD TABLE

YEAR	DURATION
5.00	36.53
19.99	47.55
25.00	61.78
50.00	72.44
00.00	83.05

WEIBULL

F(X) = PR(X(X) =	1-EXP(-(X/BETA)**ALPHA)
ALPHA=	1.156
BETA=	12.636.
MEAN=	12.007
VARIANCE=	108.437
CTR REU -	10 A17

NON-LINEAR CORRELATION IS	0.9607089
SUM SQUARE RESIDUALS IS	0.2186227
STANDARD ERROR IS	0.0801878
MEAN SQUARE DEVIATION IS	0.3858081

RETURN PERIOD TABLE

DURATION
24.96
31.64
40.15
46.40
52.52

LEAST SQUARES RESULTS - PEAK H

EXTREMAL TYPE I

F(X)=PR(X(X)=	EXP(-EXP(-(X-EPSI)/PHI))
EPSI=	374.975
PHI=	48.460
MEAN=	402.947
VARIANCE*	3862.962
STD. DEV. =	62.153

NON-LINEAR CORRELATION IS	0.9629306
SUM SQUARE RESIDUALS IS	0.2064946
STANDARD ERROR IS	0.6779318
MEAN SQUARE DEVIATION IS	8.8998112

RETURN	PERIOD	TABLE	
YEAR		PEAK	Н
5.30		478.63	
10.00		513.66	
25.00		558.90	
50.00		592.77	
100.00		626.49	

WEIBULL

F(X) = PR(X(X) =	1-EXP(-(X/DETA)++ALPHA)
ALPHA=	7.888
BETA=	426.388
MEAN=	401.273
VARIANCE=	3638.859
STD. DEV. =	60.323

NON-LINEAR CORRELATION IS	0.9038882
SUM SQUARE RESIDUALS IS	0.5192849
STANDARD ERROR IS	0.1235843
MEAN SQUARE DEVIATION IS	0.0047993

RETURN	PERIOD	TABLE	
YEAR		PEAK	Н
5.00		4.71.13	
10.00		487.88	
25.00		505.13	
50.00		515.95	
100.66		525.41	

APPENDIX D

SPSS COMMAND FILE AS APPLIED IN THE REGRESSION ANALYSIS

	rage
Command File Listing	D2
Sample Output (Excerpts)	D 3

COMMAND FILE LISTING - SPSS REGRESSION ANALYSIS

```
18$$S,T,J
284: IDENT: ROCDOPS, OPSMITH
30$: SELECT: SPSS/SPSS
48$: SYSOUT: 43, NULL
50$:LIMITS: .60K
60$: INCODE: IBMF
78RUN NAME: DURATION ANALYSIS
BOVARIABLE LIST: DUR, H. T. D
98INPUT MEDIUM: DISK
100INPUT FORMAT: FIXED (33x, F4.0, 12x, F3.0, 6x, F2.0, 6x, F3.0)
110N OF CASES: UNKNOWN
120VAR LABELS: DUR DURATION/H PEAK H/T PEAK T/D PEAK SIR/
130COMPUTE: HSQ=H++2
148COMPUTE: TSQ=T++2
150COMPUTE: STP=H/(981*TSQ)
160COMPUTE: SEV=156.13*HSQ*TSQ
178VAR LABELS: HSQ H**2/TSQ T**2/STP H OVER gT**2/SEV LH**2
180REBRESSION: VARIABLES=DUR, H, HSQ, T, TSQ/
190::REGRESSION=DUR WITH H, HSQ, T, TSQ(1) RESID=0/
210STATISTICS: ALL
226READ INPUT DATA
23@REGRESSION: VARIABLES=DUR, SEV/
248:: REGRESSION=DUR WITH SEV(1) RESID=0/
260STATISTICS: ALL
270REGRESSION: VARIABLES=DUR.D/
288::REGRESSION=DUR WITH D(1) RESID=0/
300STATISTICS: ALL
310REGRESSION: VARIABLES=DUR, STP/
320::REGRESSION=DUR WITH STP(1) RESID=0/
340STATISTICS: ALL
358SCATTERGRAM: DUR WITH H.T. HSQ.STP.SEV
360STATISTICS: ALL
370FINISH
388$: DATA: 08
390$$SELECT (P05DST50)
400$: ENDJOB
```

Sample Output (Excerpts) - SPSS Regression Analysis

DURATION ANALYSIS	1 7515								
FILE NONAM	FILE NONAME (CREATION DA	ATE = 01/16/R6)	/A6)						
•	• • • • • • • • • • • • • • • • • • • •	•	104	TIPLE	SRESSIO	• • • • •	• • • • • • • • • • • • • • • • • • • •	VARIABLE LIST	
DEPENDENT VANIANLE	WIANLE. PUR	8 T	D JRAT TON					REGRESSION CINE	-
VARTABLE (S)	VARIABLE (S) ENTERED OY STEP	NUMBER 1	429	H**2					
MULTIPLE R R SQUARE ADJUSTED R SQU STANDARD CRROR	R SQUAKE 0.45971 R SQUAKE 0.42637 ERROR 8.26466		ANALYSIS RFGRESSI RFSIOUAL	ANALYSIS OF VARIANCE Ricression Residual	of sum 1. 42.	SUM OF SQUARES 2251,38783 2868,79399	REAN SQUARE 2251, 387AS 68, 304.62	32.96099	9
	SJINDER ANIVOLES	IN THE EQU	IN THE EQUALION			VARIABLES	VARIABLES NOT IN THE EQUATION	1108	į
VARIABLE Asa (Constant)	n, nnn25 -101,43962	0£1A 0.65311	STD ERROR H 0.00004	32.961	VARIABLE H T TSØ	HETA IN -11.82259 0.17425	PARTIAL TOW	TOLERANCE F 0.00025 2. 0.87916 2.	2.744 2.051 2.051
VARIABLE (S)	VARIABLE(S) ENTERED ON STEP	NUMBER 2		T PEAK T	•	•	•	•	•
MULTIFLE R R SQUÀRE ABJUSTED R SQU STAMBARD ERROR	R SQUARE 0.48294 0.46641 R SQUARE 0.44037 ERROR A.16315		ANALYSIS Regressi Residual	AMALYSIS OF VARIANCE Regression Residual	0f SUM 2. 41.	SUN OF SQUARES 2588.06289 2752.11842	MEAN SQUARE 1194.03145 66.63705	. 4.	5
	VARIABLES	IN THE . E Q.	IN THE EQUATION		1	VARIARLES	NOT IN THE EQUI	VARIARLES NOT IN THE EQUATION	į
VARIABLE 45a T CONSTART)	8 0,00027 1,44750 -140,16131	0.72568 0.17425	STO ERROR 3 0.00005 1.31795	55.378 2.051	VARIABLE M 150	-10.45404 -0.21407	PARTIAL TO-0.22409 (-0.00000	0.00025 2.	2.115 0.000

F-LEVEL OR TOLERANCE-LEVEL INSUFFICIENT FOR FURTHER COMPUTATION

BURATION ANALYSIS

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ILE MONANE (CREATION DATE = 01/16/86)

CORRELATION COFFFICIFNTS

A VALUE OF 99,00300 IS PRINTED IF A COEFFICIENT CANNOT BE COMPUTED.

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DJRATION ANALYSIS

ILE NUNAME (CREATION DATE = 31/16/86)

VARIABLE LIST REGRESSION LIST REGRESSION RULTIPLE . PURATION **a** DEPENDENT VAKIANLF..

WARIABLE(S) ENTFRED ON STEP NUMBER 1.. SEV LH**2

3.04200 347.93038 113.62503 SUM OF SQUARES 347,93038 4772,75144 7 - 5 ANALYSIS OF VARIANCF REGRESSION RESIDUAL 0.26668 0.06795 0.04576 10.65950 MULTIPLE R R SQUARE ADJUSTED R SOUARE STANDARD FROR

----- VARIABLES NOT IN THE FOUNTION ------PARTIAL TOLERANCE SETA IN VARIABLE 3.067 VARIABLES IN THE EDUATION ---------STO ERROR 3 0.00000 AET4 0.26068 n_00000 -12_03434 (CONSTANT) VARIABLE

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MAKINUM STEP REACHED

BURATION ANALYSIS

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			SUM OF SQUARES 291.53490 4828.64692	ALE BETA IN
	. 6 R f S S			VARIAHLE
		PEAK DIR	AMALYSIS OF VARIANCE REGRESSION RESIONAL	
15/86)	•	0	27 to to	11 00 ERROR 11 OF S
ATE - '11/		NUMBER		
CRFATION O	•	FD OM STEP	0.23448 0.05694 0.05448 10.72230	20107 0
FILE MONAPE (CREATION DATE & 11/15/R6)		MEPENDENT VARIABLE BUR WARLANIE (S) ENTERED ON STEP	MULTIPLE R 0.789.77 R SQUART 0.75694 ADJUSTED R SQUARE 0.70344R STAMBARD ERROR 10.72230	VARIAHIE

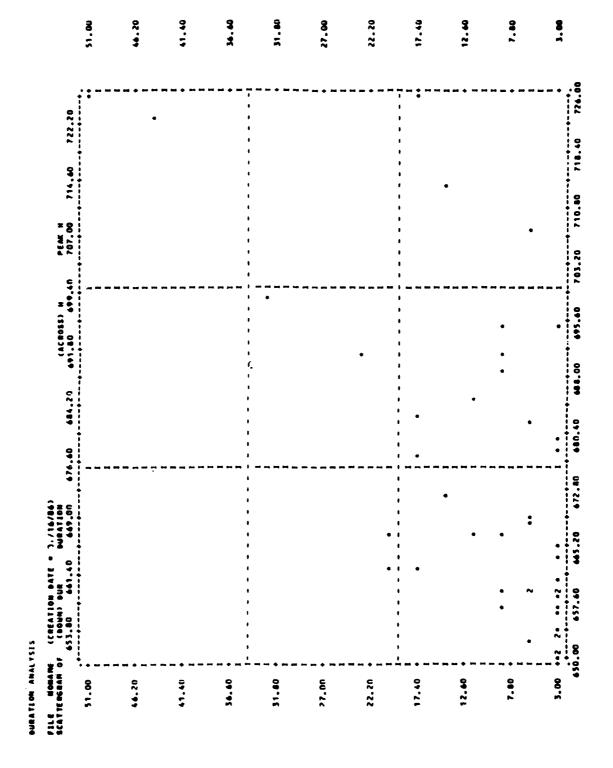
MAKINUM STIF KEACHIU

(COMSTANT)

DJRATION ANALYSIS

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	•	AHIA	1			2911	8 O 8	į	
	•	2	.E (S)	u 	<u></u>	2	2		2
111	•	DEPENDENT VARIABLE	WARTAULE(S) ENTERED ON STEP NUMBER 1	שור בונונ צ	A SEUARE	ADJUSTED R SQUAPE	STANDARD FRROR	THE REPORT OF THE PARTITION OF THE PARTIES.	VARIABLE STP
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MAKINUM STEP REACHED



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OF (BOWN) BUR BURATION
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41.40

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31.80

22.20

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916 BORANG 868778888A 07 8649	(CREATION DATE = 31/16/96) (BOWN) DUA BURATION 5440.009396334.889112248.000	71 / 14 / 84) Buration 1522 ra_oda2 68 19	ME (CEMENT) BUTE - 31/14/20)	(ACR655) SEV 0.001575984.80	LH++2 2011808.002447712.	00.863416.00
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